

Chapter 8

Note to the Instructor: Some of the revised control charts provided in the solutions retain the removed points only as place holders. Other revised control charts have been created by removing the out-of-control point from the worksheet and constructing the control charts again.

Section 8-4

8-1. a)

\bar{x} chart

$$UCL = CL + A_2\bar{r} = 39.42$$

$$CL = 33.95$$

$$LCL = CL - A_2\bar{r} = 28.48$$

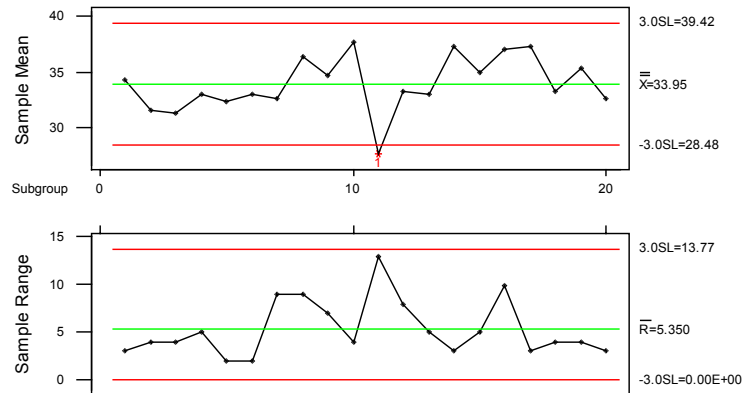
R chart

$$UCL = D_4\bar{r} = 13.77$$

$$CL = 5.35$$

$$LCL = D_3\bar{r} = 0(5.35) = 0$$

Xbar/R Chart for x1-x3



b)

\bar{x} chart

$$UCL = 39.34$$

$$CL = 34.28$$

$$LCL = 29.22$$

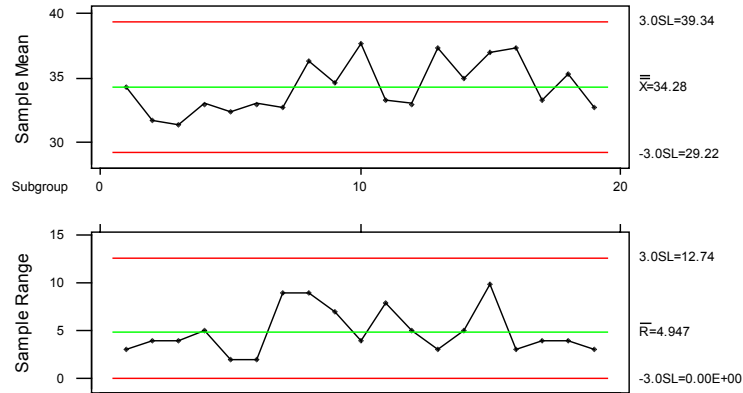
R chart

$$UCL = 12.74$$

$$CL = 4.947$$

$$LCL = D_3\bar{r} = 0(4.947) = 0$$

Xbar/R Chart for x1-x3



8-2. a) $\bar{\bar{x}} = \frac{378.50}{20} = 18.925$ $\bar{r} = \frac{7.80}{20} = 0.39$

\bar{x} chart

R chart

$UCL = CL + A_2\bar{r} = 18.925 + 0.729(0.39) = 19.21$

$UCL = D_4\bar{r} = 2.282(0.39) = 0.89$

$CL = 18.925$

$CL = 0.39$

$LCL = CL - A_2\bar{r} = 18.925 - 0.729(0.39) = 18.64$

$LCL = D_3\bar{r} = 0(0.39) = 0$

b)

$\hat{\mu} = \bar{\bar{x}} = 18.925$

$\hat{\sigma} = \frac{\bar{r}}{d_2} = \frac{0.39}{2.059} = 0.1894$

8-3. a) The trial control limits are as follows.

\bar{x} chart

R chart

$UCL = 18.20$

$UCL = 9.581$

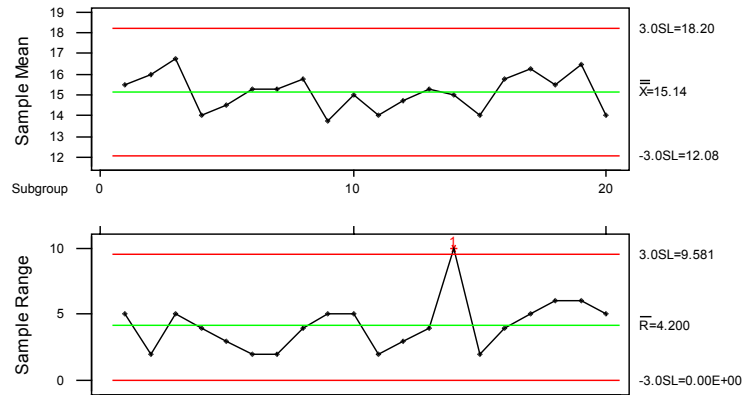
$CL = 15.14$

$CL = 4.2$

$LCL = 12.08$

$LCL = 0$

Xbar/R Chart for Problem 8-3

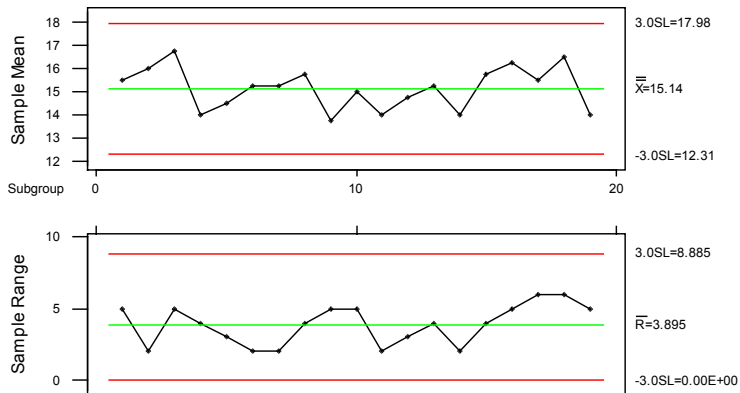


Based on the control charts, there is a single observation beyond the control limits. Observation 14 is above the upper control limit on the R chart.

b) With Observation 14 removed, the control limits and charts are as follows.

| \bar{x} chart | R chart |
|-----------------|-------------|
| UCL = 17.98 | UCL = 8.885 |
| CL = 15.14 | CL = 3.895 |
| LCL = 12.31 | LCL = 0 |

Previous Xbar/R Chart for 8-3



All points are within the control limits. The process is said to be in statistical control.

8-4. $\bar{\bar{x}} = 20.0$ $\frac{\bar{r}}{d_2} = 1.4$ $d_2 = 2.534$ giving $\bar{r} = 1.4(2.534) = 3.5476$

| \bar{x} chart | R chart |
|--|--|
| UCL = CL + A ₂ \bar{r} = 20.0 + 0.483(3.5476) = 21.71 | UCL = D ₄ \bar{r} = 2.004(3.53476) = 7.11 |
| CL = 20.0 | CL = 3.5476 |
| LCL = CL - A ₂ \bar{r} = 20.0 - 0.483(3.5476) = 18.29 | LCL = D ₃ \bar{r} = 0(3.5476) = 0 |

8-5. a) $\bar{\bar{x}} = \frac{7657}{25} = 306.28$ $\bar{r} = \frac{1180}{25} = 47.2$

\bar{x} chart

R chart

$UCL = CL + A_2\bar{r} = 306.28 + 0.729(47.2) = 340.69$

$UCL = D_4\bar{r} = 2.282(47.2) = 107.71$

$CL = 223$

$CL = 47.2$

$LCL = CL - A_2\bar{r} = 306.28 - 0.729(47.2) = 271.87$

$LCL = D_3\bar{r} = 0(47.2) = 0$

b)

$\hat{\mu} = \bar{\bar{x}} = 306.28$

$\hat{\sigma} = \frac{\bar{r}}{d_2} = \frac{47.2}{2.059} = 22.92$

8-6. a)

\bar{x} chart

R chart

$UCL = 0.06352$

$UCL = 0.002106$

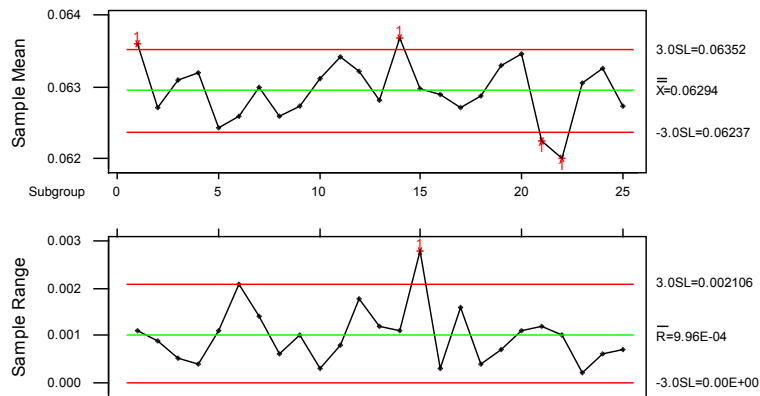
$CL = 0.06294$

$CL = 0.000996$

$LCL = 0.06237$

$LCL = 0$

Xbar/R Chart for x1-x5



There are several points out of control. The control limits need to be revised. The points are 1, 14, 21, 22; or outside the control limits of the R chart: 15

b) Second revision, observations 5 and 6 are removed. Observation 12 is removed after the third revision. Observations 17 and 20 are removed after the fourth revision. Observations 8 and 11 are removed after the fifth revision. Observation 7 is removed after the sixth revision. After several revisions, the control chart is finally in control.

\bar{x} chart

R chart

$UCL = 0.06332$

$UCL = 0.001269$

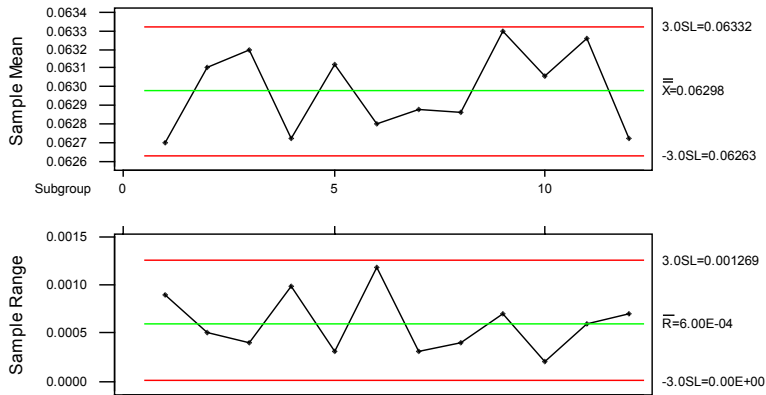
$CL = 0.06298$

$CL = 0.0006$

$LCL = 0.06263$

$LCL = 0$

Xbar/R Chart for x1-x5



8-7.

a)

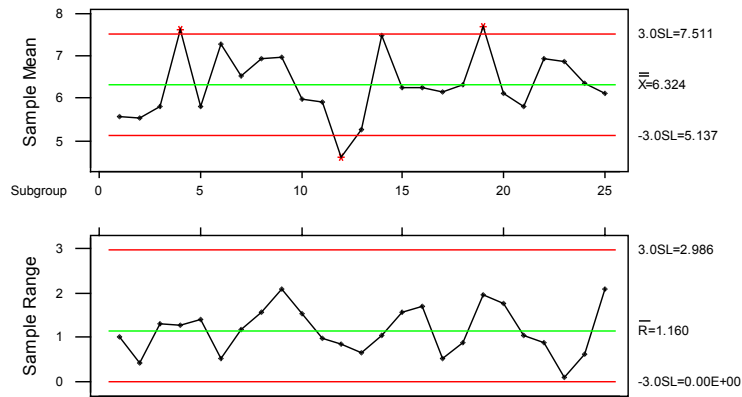
\bar{x} chart

UCL = 7.511
 CL = 6.324
 LCL = 5.137

R chart

UCL = 2.986
 CL = 1.160
 LCL = 0

Xbar/R Chart for 8-7



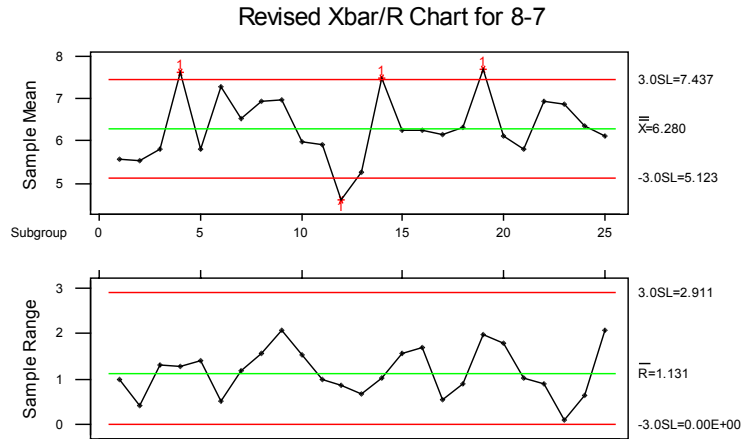
b)

\bar{x} chart

UCL = 7.437
 CL = 6.280
 LCL = 5.123

R chart

UCL = 2.911
 CL = 1.131
 LCL = 0



One point still beyond the control limits. Revise the limits again.

\bar{x} chart

R chart

UCL = 7.385

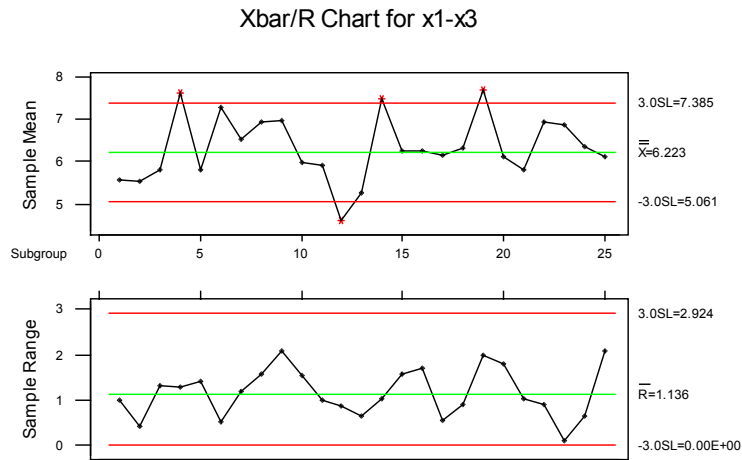
UCL = 2.924

CL = 6.223

CL = 1.136

LCL = 5.061

LCL = 0



No point lies beyond the control limits, the process is now in control.

Section 8-5

8-8. a)

\bar{x} chart

\overline{MR} chart

UCL = 63.25

UCL = 12.04

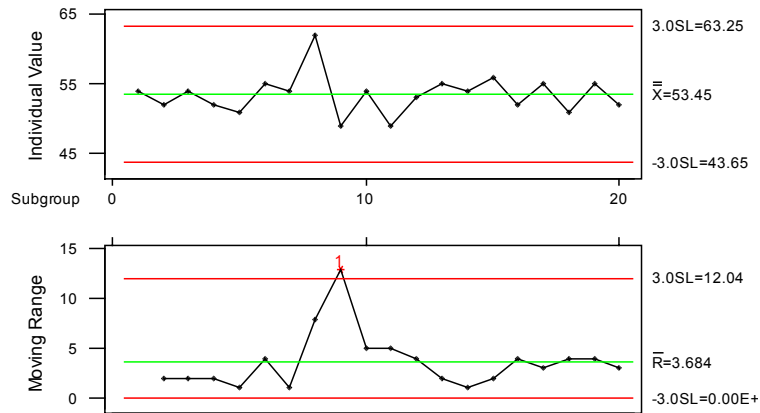
CL = 53.45

CL = 3.684

LCL = 43.65

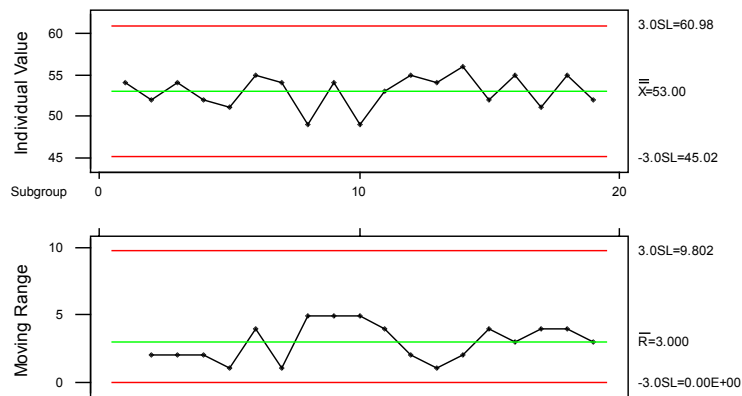
LCL = 0

I and MR Chart for Hardness Problem 8-8



There is a single observation beyond the upper control limit on the R chart. Remove this observation (#8 in the dataset) and revise the control limits.

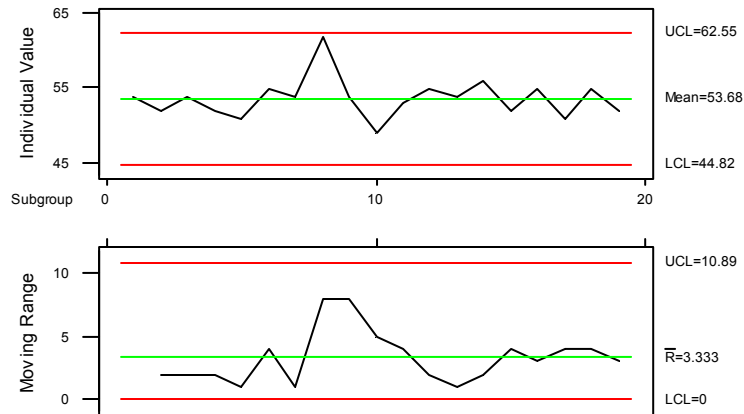
I and MR Chart for Hardness



The process now appears to be in control.

There is a single observation beyond the upper control limit on the R chart. Remove this observation (#9 in the dataset) and revise the control limits.

I and MR Chart for Hardness_1



b)

$$\hat{\mu} = \bar{x} = 53$$

$$\hat{\sigma} = \frac{\overline{MR}}{d_2} = \frac{3.00}{1.128} = 2.66$$

if removing # 9 then

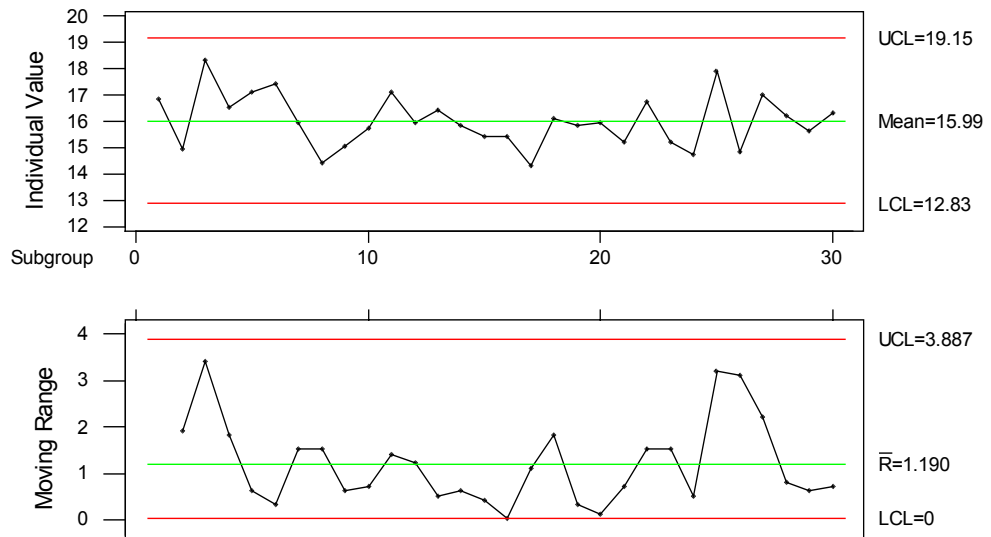
$$\hat{\mu} = \bar{x} = 53.68$$

$$\hat{\sigma} = \frac{\overline{MR}}{d_2} = \frac{3.33}{1.128} = 2.952$$

8-9. a) X chart: UCL = 19.15, LCL = 12.83, Mean = 15.99

MR chart: UCL = 3.887, LCL = 0, Mean = 1.190

I and MR Chart for x



There are no points beyond the control limits. The process appears to be in control.

b) Estimates are: $\hat{\mu} = 15.99$, $\hat{\sigma} = 1.19/1.128 = 1.055$

8-10. a)

x chart

UCL = 34.63

CL = 17.63

LCL = 1.952

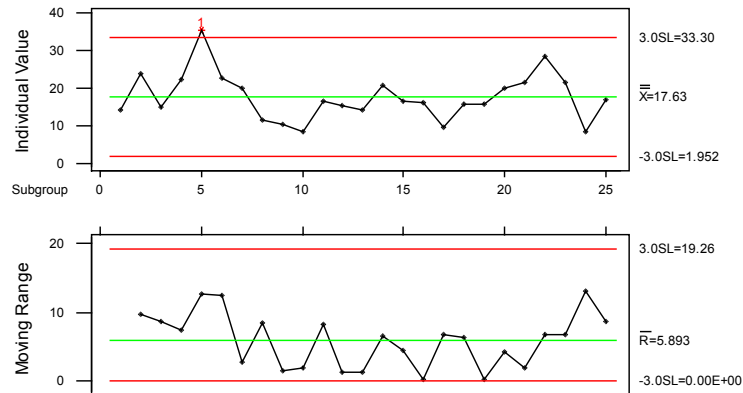
MR chart

UCL = 19.26

CL = 5.893

LCL = 0

I and MR Chart for Diameter



Observation 5 is beyond the upper control limit on the individual chart. Revise the control limits.

x chart

UCL = 30.35

CL = 16.89

LCL = 3.431

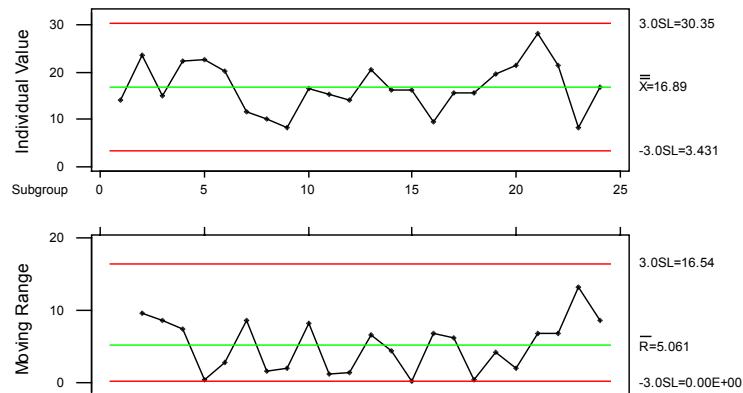
MR chart

UCL = 16.54

CL = 5.061

LCL = 0

I and MR Chart for Diameter



All observations are within the control limits.

b)

$$\hat{\mu} = \bar{x} = 16.89$$

$$\hat{\sigma} = \frac{\overline{mr}}{d_2} = \frac{5.061}{1.128} = 4.487$$

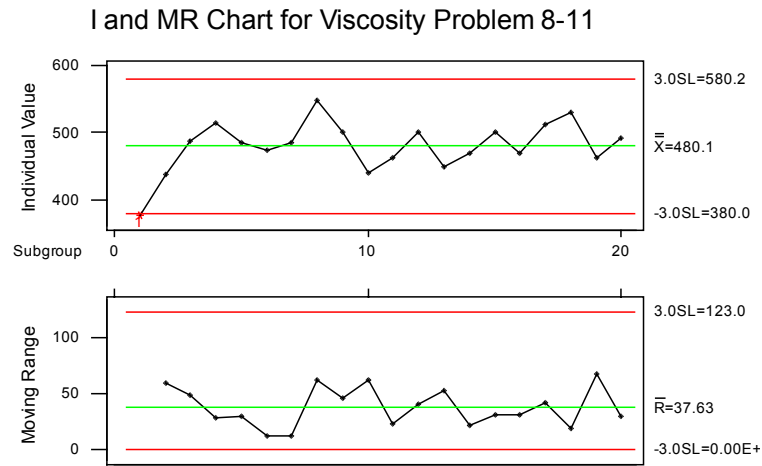
8-11. a)

x chart

UCL = 580.2
CL = 480.1
LCL = 380

\overline{MR} chart

UCL = 123
CL = 37.63
LCL = 0



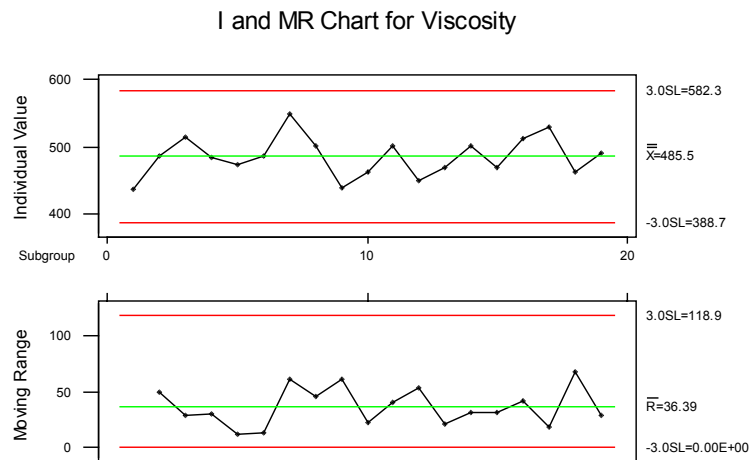
The first observation is below the lower control limit. Revise the control limits.

x chart

UCL = 582.3
CL = 485.5
LCL = 388.7

\overline{MR} chart

UCL = 118.9
CL = 36.39
LCL = 0



The process is in control.

$$b) \hat{\mu} = \bar{x} = 485.5 \quad \hat{\sigma} = \frac{\overline{mr}}{d_2} = \frac{36.39}{1.128} = 32.26$$

Section 8-6

- 8-12. a) If the process uses 66.7% of the specification band, then $6\sigma = 0.667(USL - LSL)$ then assume $\bar{\bar{x}} = \mu$ since the process is centered
 $3\sigma = 0.667(USL - \bar{\bar{x}}) = 0.667(\bar{\bar{x}} - LSL) = 0.667(USL - \mu)$
 $4.5\sigma = USL - \mu = LSU - \mu$

$$C_{pk} = \min\left[\frac{4.5\sigma}{3\sigma}, \frac{4.5\sigma}{3\sigma}\right] = 1.5$$

Since C_p and C_{pk} exceeds unity, the natural tolerance limits lie inside the specification limits and very few defective units will be produced.

- b) Assuming a normal distribution with $6\sigma = 0.667(USL - LSL)$ and a centered process, then $3\sigma = 0.667(USL - \mu)$. Consequently, $USL - \mu = 4.5\sigma$ and $\mu - LSL = 4.5\sigma$

$$\begin{aligned} P(X > USL) &= P\left(Z > \frac{4.5\sigma}{\sigma}\right) \\ &= P(Z > 4.5) \\ &= 1 - P(Z < 4.5) \\ &= 1 - 1 \\ &= 0 \end{aligned}$$

By symmetry, the fraction defective is $2[P(X > USL)] = 0$.

8-13. a) $\hat{\sigma} = \frac{\bar{r}}{d_2} = \frac{4.947}{1.693} = 2.922$ or $\hat{\sigma} = 0.0002922$

$$C_p = \frac{USL - LSL}{6(\hat{\sigma})} = \frac{0.4040 - 0.4020}{6(0.0002922)} = 1.141$$

$$\begin{aligned} C_{PK} &= \min\left[\frac{USL - \bar{\bar{x}}}{3\hat{\sigma}}, \frac{\bar{\bar{x}} - LSL}{3\hat{\sigma}}\right] \\ &= \min\left[\frac{0.4040 - 0.40348}{3(0.0002922)}, \frac{0.40348 - 0.4020}{3(0.0002922)}\right] \\ &= \min[0.5932, 1.688] \\ &= 0.5932 \end{aligned}$$

Since C_p exceeds unity, the natural tolerance limits lie inside the specification limits and very few defective units will be produced.

Since $C_{pk} \neq C_p$ the process is off center.

- b) Assuming a normal distribution with $\hat{\mu} = 0.40348$ and $\hat{\sigma} = 0.0002922$

$$\begin{aligned}
P(X < LSL) &= P\left(Z < \frac{LSL - \hat{\mu}}{\hat{\sigma}}\right) \\
&= P(Z < -5.07) \\
&= 0
\end{aligned}$$

$$\begin{aligned}
P(X > USL) &= P\left(Z > \frac{USL - \hat{\mu}}{\hat{\sigma}}\right) \\
&= P(Z > 1.78) \\
&= 1 - P(Z < 1.78) \\
&= 1 - 0.9624 \\
&= 0.0376
\end{aligned}$$

Therefore, the proportion nonconforming is given by
 $P(X < LSL) + P(X > USL) = 0.0376 + 0 = 0.0376$

- 8-14. a) Assuming a normal distribution with $\hat{\mu} = 18.925$ and $\hat{\sigma} = \frac{\bar{r}}{d_2} = \frac{0.39}{2.059} = 0.189$

$$\begin{aligned}
P(X < LSL) &= P\left(Z < \frac{LSL - \hat{\mu}}{\hat{\sigma}}\right) \\
&= P\left(Z < \frac{18.00 - 18.925}{0.189}\right) \\
&= P(Z < -4.89) \\
&= 0
\end{aligned}$$

$$\begin{aligned}
P(X > USL) &= P\left(Z > \frac{USL - \hat{\mu}}{\hat{\sigma}}\right) \\
&= P\left(Z > \frac{19.00 - 18.925}{0.189}\right) \\
&= P(Z > 0.40) \\
&= 1 - P(Z < 0.40) \\
&= 1 - 0.65542 \\
&= 0.34458
\end{aligned}$$

Therefore, the proportion nonconforming is given by
 $P(X < LSL) + P(X > USL) = 0 + 0.3446 = 0.3446$

b)

$$C_p = \frac{USL - LSL}{6(\hat{\sigma})} = \frac{19.00 - 18.00}{6(0.189)} = 0.882$$

$$\begin{aligned}
C_{pk} &= \min\left[\frac{USL - \bar{\bar{x}}}{3\hat{\sigma}}, \frac{\bar{\bar{x}} - LSL}{3\hat{\sigma}}\right] \\
&= \min\left[\frac{19.00 - 18.925}{3(0.189)}, \frac{18.925 - 18.00}{3(0.189)}\right] \\
&= \min[0.132, 1.63] \\
&= 0.132
\end{aligned}$$

Since C_p less than unity, many defective units are being produced.

$C_{pk} \neq C_p$ the process is not centered.

8-15. $\hat{\sigma} = \frac{\bar{r}}{d_2} = \frac{3.895}{2.059} = 1.892$ or 0.01892 mm

$$\begin{aligned} \bar{x} &= 15.14 \\ P(X > 18) + P(X < 12) \\ &= P\left(Z > \frac{18 - 15.14}{1.892}\right) + P\left(Z < \frac{12 - 15.14}{1.892}\right) \\ &= P(Z > 1.51) + P(Z < -1.66) \\ &= 0.065522 + 0.048457 \\ &= 0.113979 \end{aligned}$$

$$Cp = \frac{18 - 12}{6(1.892)} = 0.5285$$

With the Cp less than unity, the process capability appears to be poor.

8-16. a) If the process uses 85% of the spec band then $6\sigma = 0.85(\text{USL} - \text{LSL})$ and

$$Cp = \frac{\text{USL} - \text{LSL}}{0.85(\text{USL} - \text{LSL})} = \frac{1}{0.85} = 1.18$$

Assume $\bar{x} = \mu$ and $3\sigma = 0.85(\text{USL} - \bar{x}) = 0.85(\mu - \text{LSL})$

Therefore,

$$Cpk = \min\left[\frac{3.53\hat{\sigma}}{3\hat{\sigma}}, \frac{3.53\hat{\sigma}}{3\hat{\sigma}}\right] = 1.18$$

Since Cp and Cpk exceed unity, the natural tolerance limits lie inside the specification limits and very few defective units will be produced.

b) Assuming a normal distribution with $6\sigma = 0.85(\text{USL} - \text{LSL})$ and a centered process, then $3\sigma = 0.85(\text{USL} - \mu)$. Consequently, $\text{USL} - \mu = 3.53\sigma$ and $\mu - \text{LSL} = 3.53\sigma$

$$\begin{aligned} P(X > \text{USL}) &= P\left(Z > \frac{3.53\sigma}{\sigma}\right) \\ &= P(Z > 3.53) \\ &= 1 - P(Z < 3.53) \\ &= 1 - 0.9998 \\ &= 0.0002 \end{aligned}$$

By symmetry, the fraction defective is $2[P(X > \text{USL})] = 0.0004$.

8-17. Assuming a normal distribution with $\hat{\mu} = 306.28$ and $\hat{\sigma} = 22.923$

$$\begin{aligned}
P(X < LSL) &= P\left(Z < \frac{LSL - \hat{\mu}}{\hat{\sigma}}\right) \\
&= P\left(Z < \frac{260 - 306.28}{22.923}\right) \\
&= P(Z < -2.02) \\
&= 0.0217
\end{aligned}$$

$$\begin{aligned}
P(X > USL) &= P\left(Z > \frac{USL - \hat{\mu}}{\hat{\sigma}}\right) \\
&= P\left(Z > \frac{340 - 306.28}{22.923}\right) \\
&= P(Z > 1.47) \\
&= 1 - P(Z < 1.47) \\
&= 1 - 0.9292 \\
&= 0.0708
\end{aligned}$$

Therefore, the proportion nonconforming is given by
 $P(X < LSL) + P(X > USL) = 0.0217 + 0.0708$
 $= 0.0925$

$$Cp = \frac{USL - LSL}{6(\hat{\sigma})} = \frac{340 - 260}{6(22.923)} = 0.582$$

$$\begin{aligned}
Cpk &= \min\left[\frac{USL - \bar{\bar{x}}}{3\hat{\sigma}}, \frac{\bar{\bar{x}} - LSL}{3\hat{\sigma}}\right] \\
&= \min\left[\frac{340 - 306.28}{3(22.923)}, \frac{306.28 - 260}{3(22.923)}\right] \\
&= \min[0.490, 0.673] \\
&= 0.490
\end{aligned}$$

The process capability is marginal.

8-18. Assuming a normal distribution with $\hat{\mu} = 20.0$ and $\hat{\sigma} = 1.4$

$$Cp = \frac{USL - LSL}{6(\hat{\sigma})} = \frac{25 - 15}{6(1.4)} = 1.19$$

$$\begin{aligned}
Cpk &= \min\left[\frac{USL - \bar{\bar{x}}}{3\hat{\sigma}}, \frac{\bar{\bar{x}} - LSL}{3\hat{\sigma}}\right] \\
&= \min\left[\frac{25 - 20}{3(1.4)}, \frac{20 - 15}{3(1.4)}\right] \\
&= \min[1.19, 1.19] \\
&= 1.19
\end{aligned}$$

The process is capable.

8-19. Assuming a normal distribution with $\hat{\mu} = 6.223$ and $\hat{\sigma} = \frac{1.136}{1.693} = 0.671$

$$Cp = \frac{USL - LSL}{6(\hat{\sigma})} = \frac{6.5 - 5.5}{6(0.671)} = 0.248$$

$$\begin{aligned} Cpk &= \min \left[\frac{USL - \bar{\bar{x}}}{3\hat{\sigma}}, \frac{\bar{\bar{x}} - LSL}{3\hat{\sigma}} \right] \\ &= \min \left[\frac{6.5 - 6.223}{3(0.671)}, \frac{6.223 - 5.5}{3(0.671)} \right] \\ &= \min [0.138, 0.359] \\ &= 0.138 \end{aligned}$$

The process capability is poor.

8-20. Assuming a normal distribution with $\hat{\mu} = 0.06298$ and $\hat{\sigma} = \frac{0.0006}{2.326} = 0.00026$

The natural tolerance limits are then

$$\begin{aligned} \hat{\mu} \pm 3\hat{\sigma} &= 0.06298 \pm 3(0.00026) \\ &= (0.0622, 0.0638) \end{aligned}$$

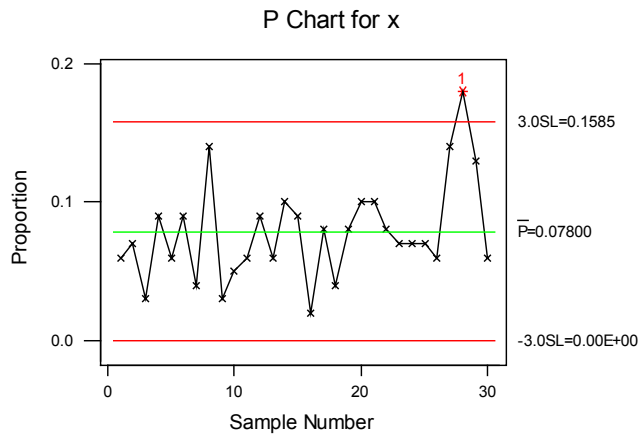
8-21. Assuming a normal distribution with $\hat{\mu} = 485.5$ and $\hat{\sigma} = 32.26$

The natural tolerance limits are then

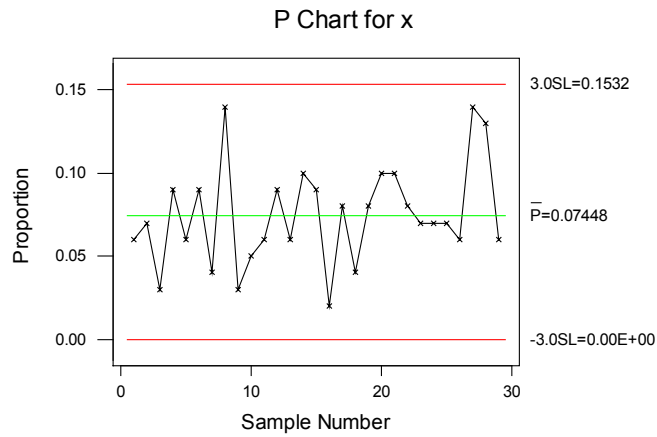
$$\begin{aligned} \hat{\mu} \pm 3\hat{\sigma} &= 485.5 \pm 3(32.26) \\ &= (388.72, 582.28) \end{aligned}$$

Section 8-7

8-22. a) The control limits are
 UCL = 0.1585
 CL = 0.078
 LCL = 0

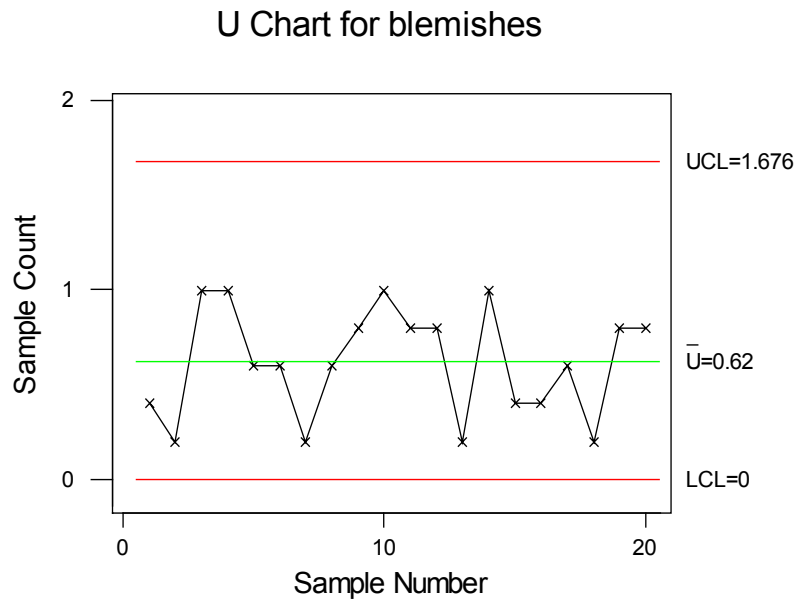


b) The process is out of control, revised limits are
 UCL = 0.1532
 CL = 0.7448
 LCL = 0



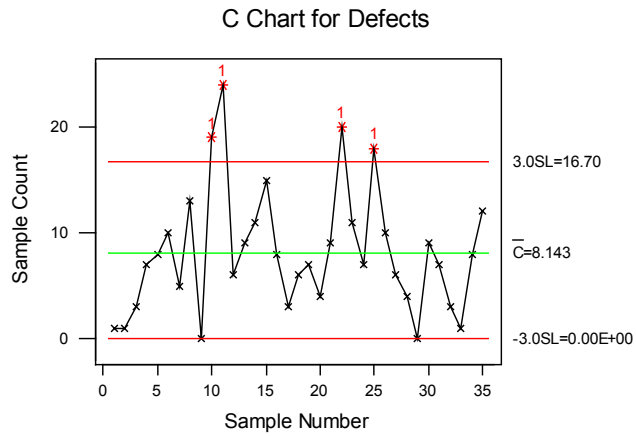
The process is now in control.

- 8-23. a) The control limits are
 UCL = 1.676
 CL = 0.62
 LCL = 0



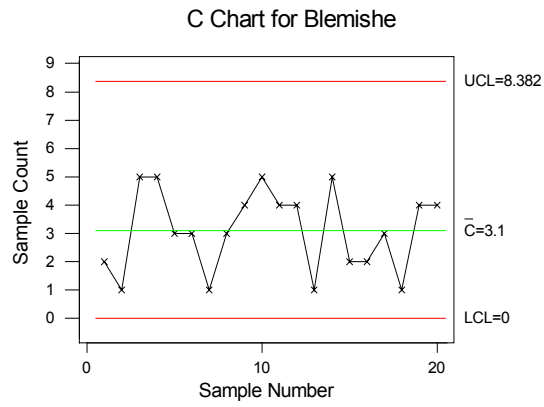
b) The process appears to be in control

- 8-24. The control limits are
 UCL = 16.70
 CL = 8.143
 LCL = 0



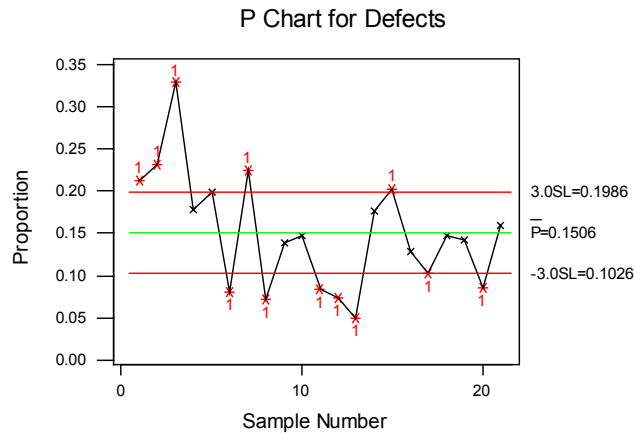
The process is not in control. This is evident by the fact that there are several points beyond the control limits.

8-25.



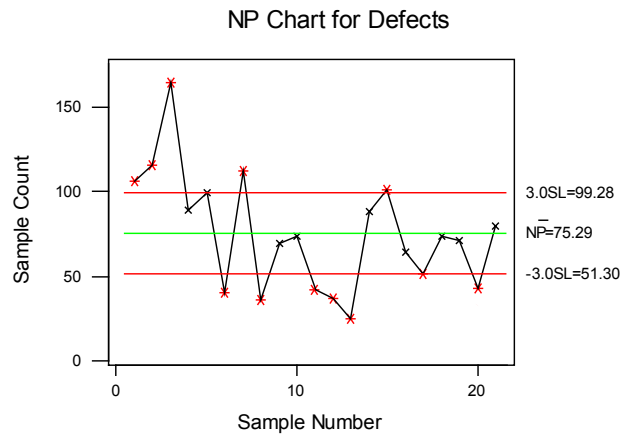
The U chart and the C chart both showed that the process is in control.

8-26. a) P-chart



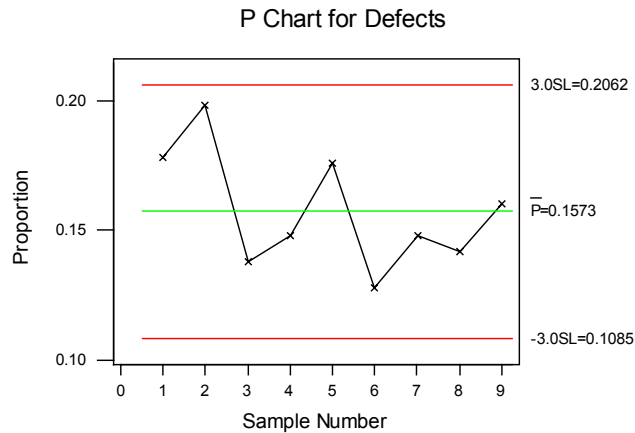
The samples with out-of-control points are 1, 2, 3, 6, 7, 8, 11, 12, 13, 15, 17, 20. There are several points are out of control. The control limits need to be revised.

nP chart



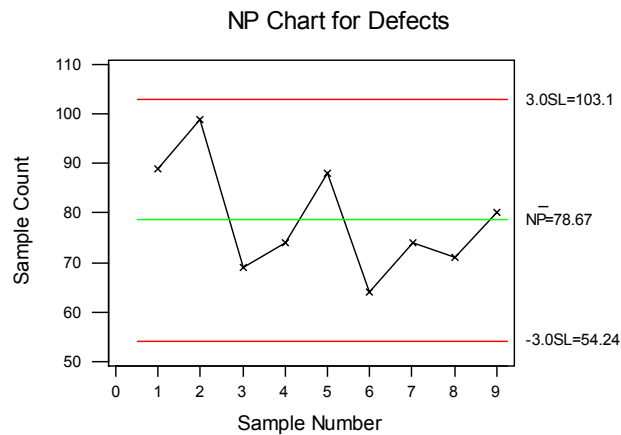
The samples with out-of-control points are 1, 2, 3, 6, 7, 8, 11, 12, 13, 15, 17, 20. There are several points are out of control. The control limits need to be revised.

b) P – chart revised



There are no points out of control for the revised limits.

nP chart - revised



The process is now in control.

Section 8-8

8-27. a) $\bar{x} = 74.01$ $\sigma_{\bar{x}} = 0.0045$ $\mu = 74.01$

$$\begin{aligned}
 & P(73.9865 < \bar{X} < 74.0135) \\
 &= P\left(\frac{73.9865 - 74.01}{0.0045} < \frac{X - \mu}{\hat{\sigma}_{\bar{x}}} < \frac{74.0135 - 74.01}{0.0045}\right) \\
 &= P(-5.22 < Z < 0.78) \\
 &= P(Z < 0.78) - P(Z < -5.22) \\
 &= P(Z < 0.78) - [1 - P(Z < 5.22)] \\
 &= 0.7823 - (1 - 1) \\
 &= 0.7823
 \end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.7823 = 0.2177$.

$$b) ARL = \frac{1}{p} = \frac{1}{0.2177} = 4.6$$

$$8-28. a) \mu + 3 \frac{\sigma}{\sqrt{n}} = UCL$$

$$100 + 3 \frac{\sigma}{\sqrt{6}} = 106$$

$$\sigma = \frac{\sqrt{6}}{3} (106 - 100) = 4.9$$

$$b) \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{4.9}{2.45} = 2, \quad \mu = 105$$

$$\begin{aligned} & P(94 < \bar{X} < 106) \\ &= P\left(\frac{94 - 105}{2} < \frac{\bar{X} - \mu}{\hat{\sigma}_{\bar{x}}} < \frac{106 - 105}{2}\right) \\ &= P(-5.5 < Z < 0.5) \\ &= P(Z < 0.5) - P(Z < -5.5) \\ &= P(Z < 0.5) - [1 - P(Z < 5.5)] \\ &= 0.69146 - 0 \\ &= 0.69146 \end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.69146 = 0.30854$.

$$c) ARL = \frac{1}{p} = \frac{1}{0.30854} = 3.24$$

$$8-29. a) \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{2.922}{\sqrt{3}} = 1.69, \quad \mu = 38$$

$$\begin{aligned} & P(29.22 < \bar{X} < 39.34) \\ &= P\left(\frac{29.22 - 39}{1.69} < \frac{\bar{X} - \mu}{\hat{\sigma}_{\bar{x}}} < \frac{39.34 - 39}{1.69}\right) \\ &= P(-5.78 < Z < 0.201) \\ &= P(Z < 0.201) - P(Z < -5.78) \\ &= 0.5797 - 0 \\ &= 0.5797 \end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.5797 = 0.4203$.

$$b) ARL = \frac{1}{p} = \frac{1}{0.4203} = 2.38$$

$$8-30. a) \hat{\sigma} = \frac{\bar{R}}{d_2} = \frac{0.39}{2.059} = 0.189 \quad \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{0.189}{\sqrt{4}} = 0.0945, \quad \mu = 18.7$$

$$\begin{aligned}
& P(18.641 < X < 19.209) \\
& = P\left(\frac{18.641 - 19.1}{0.0945} < \frac{X - \mu}{\sigma_x} < \frac{19.209 - 19.1}{0.0945}\right) \\
& = P(-4.86 < Z < 1.15) \\
& = P(Z < 1.15) - P(Z < -4.86) \\
& = 0.874928 - 0 \\
& = 0.874928
\end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.874928 = 0.125072$

$$b) ARL = \frac{1}{p} = \frac{1}{0.125072} = 7.995$$

$$8-31. \quad a) \hat{\sigma} = \frac{\bar{R}}{d_2} = \frac{3.895}{2.059} = 1.89 \quad \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{1.89}{\sqrt{4}} = 0.945, \mu = 14.5$$

$$\begin{aligned}
& P(12.31 < \bar{X} < 17.98) \\
& = P\left(\frac{12.31 - 12.8}{0.945} < \frac{\bar{X} - \mu}{\sigma_{\bar{x}}} < \frac{17.89 - 12.8}{0.945}\right) \\
& = P(-0.519 < Z < 5.39) \\
& = P(Z < 5.39) - P(Z < -0.518) \\
& = 1 - 0.3022 \\
& = 0.6978
\end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.6978 = 0.3022$.

$$b) ARL = \frac{1}{p} = \frac{1}{0.3022} = 3.31$$

$$8-32. \quad a) \hat{\sigma} = \frac{\bar{R}}{d_2} = 1.4 \quad \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{1.4}{\sqrt{6}} = 0.572, \mu = 19$$

$$\begin{aligned}
& P(18.12 < X < 21.88) \\
& = P\left(\frac{18.29 - 18.5}{0.572} < \frac{\bar{X} - \mu}{\sigma_{\bar{x}}} < \frac{21.71 - 18.5}{0.572}\right) \\
& = P(-0.37 < Z < 5.61) \\
& = P(Z < 5.61) - P(Z < -0.37) \\
& = 1 - 0.355691 \\
& = 0.644309
\end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.644309 = 0.355691$.

$$b) ARL = \frac{1}{p} = \frac{1}{0.355691} = 2.811$$

$$8-33. \quad a) \hat{\sigma} = \frac{47.2}{2.059} = 22.92 \quad \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{22.92}{\sqrt{4}} = 11.46$$

$$\begin{aligned}
& P(271.87 < \bar{X} < 340.69) \\
& = P\left(\frac{271.87 - 310}{11.46} < \frac{\bar{X} - \mu}{\sigma_{\bar{x}}} < \frac{340.69 - 310}{11.46}\right) \\
& = P(-3.33 < Z < 2.68) \\
& = P(Z < 2.68) - P(Z < -3.33) \\
& = 0.9963 - 0.00043 \\
& = 0.99587
\end{aligned}$$

The probability that this shift will be detected on the next sample is $= 1 - 0.99587 = 0.00413$.

$$b) ARL = \frac{1}{p} = \frac{1}{0.00413} = 242.13$$

$$8-34. a) \hat{\sigma} = 0.00024 \quad \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{0.00024}{\sqrt{5}} = 0.000107, \mu = 0.0628$$

$$\begin{aligned}
& P(0.06266 < \bar{X} < 0.06331) \\
& = P\left(\frac{0.06266 - 0.063}{0.000107} < \frac{\bar{X} - \mu}{\sigma_{\bar{x}}} < \frac{0.06331 - 0.063}{0.000107}\right) \\
& = P(-3.18 < Z < 2.90) \\
& = P(Z < 2.90) - P(Z < -3.18) \\
& = 0.998134 - 0.000736 \\
& = 0.997398
\end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.997398 = 0.002602$

$$b) ARL = \frac{1}{p} = \frac{1}{0.002602} = 384.32$$

$$8-35. a) \hat{\sigma} = \frac{1.136}{1.693} = 0.671 \quad \hat{\sigma}_{\bar{x}} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{0.671}{\sqrt{3}} = 0.387, \mu = 5.5$$

$$\begin{aligned}
& P(5.061 < \bar{X} < 7.385) \\
& = P\left(\frac{5.061 - 6.8}{0.387} < \frac{\bar{X} - \mu}{\sigma_{\bar{x}}} < \frac{7.385 - 6.8}{0.387}\right) \\
& = P(-4.49 < Z < 1.51) \\
& = P(Z < 1.51) - P(Z < -4.49) \\
& = 0.93448 - 0 \\
& = 0.93448
\end{aligned}$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.93448 = 0.06552$.

$$b) ARL = \frac{1}{p} = \frac{1}{0.06552} = 15.26$$

Section 8-9

8-36. a) One-Way ANOVA Table

| Source | DF | SS | MS | F | P |
|--------|----|--------|---------|---------|---------|
| Wafer | 19 | 7209.4 | 379.442 | 5.01024 | 0.00001 |

| | | | |
|---------------|----|---------|--------|
| Repeatability | 40 | 3029.3 | 75.733 |
| Total | 59 | 10238.7 | |

There is significant difference in the parts used in the study based on these three measurements.

b) Gage R&R

| Source | VarComp | %Contribution (of VarComp) |
|------------------|---------|-------------------------------|
| Total Gage R&R | 75.73 | 42.79 |
| Repeatability | 75.73 | 42.79 |
| Wafer -to- Wafer | 101.24 | 57.21 |
| Total Variation | 176.97 | 100.00 |

| c) Source | StdDev (SD) | Study Var (5.15*SD) | %Study Var (%SV) |
|-----------------|----------------|------------------------|---------------------|
| Total Gage R&R | 8.7025 | 44.8178 | 65.42 |
| Repeatability | 8.7025 | 44.8178 | 65.42 |
| Part-to-Part | 10.0616 | 51.8174 | 75.63 |
| Total Variation | 13.3030 | 68.5104 | 100.00 |

The gauge contributes about 65.42% of total variability.

8-37. a) Analysis of Variance for purity

| Source | DF | SS | MS | F | P |
|---------|----|---------|---------|------|-------|
| Measure | 9 | 0.32050 | 0.03561 | 0.43 | 0.889 |
| Error | 10 | 0.82500 | 0.08250 | | |
| Total | 19 | 1.14550 | | | |

There is no significant difference in the measuring device.

b) Source Variance Error Expected Mean Square for Each Term
 component term (using unrestricted model)

1 measure -0.02344 2 (2) + 2(1)

2 Error 0.08250 (2)

total = 0 + 0.08250 = 0.08250

c) 100%. You have a desirable situation.

8-38. a) One-Way ANOVA Table

| Source | DF | SS | MS | F | P |
|---------------|----|---------|---------|---------|---------|
| Unit | 7 | 306.167 | 43.7382 | 1.93137 | 0.18804 |
| Repeatability | 8 | 181.170 | 22.6463 | | |
| Total | 15 | 487.337 | | | |

There is no significant difference in the units used in the study.

b) Gage R&R

| Source | VarComp | %Contribution (of VarComp) |
|-----------------|---------|-------------------------------|
| Total Gage R&R | 22.646 | 68.23 |
| Repeatability | 22.646 | 68.23 |
| Unit-to-Unit | 10.546 | 31.77 |
| Total Variation | 33.192 | 100.00 |

c)

| Source | StdDev (SD) | Study Var (5.15*SD) | %Study Var (%SV) |
|-----------------|----------------|------------------------|---------------------|
| Total Gage R&R | 4.75881 | 24.5079 | 82.60 |
| Repeatability | 4.75881 | 24.5079 | 82.60 |
| Part-to-Part | 3.24746 | 16.7244 | 56.37 |
| Total Variation | 5.76127 | 29.6705 | 100.00 |

The gauge contributes about 82.6% of total variability.

8-39. a) Analysis of Variance for Strength

| Source | DF | SS | MS | F | P |
|---------------|----|--------|-------|------|-------|
| Part | 5 | 457.8 | 91.6 | 1.61 | 0.307 |
| Operator | 1 | 121.4 | 121.4 | 2.14 | 0.204 |
| Part*Operator | 5 | 284.1 | 56.8 | 0.46 | 0.799 |
| Error | 12 | 1485.8 | 123.8 | | |
| Total | 23 | 2349.1 | | | |

None of the terms is significant.

| b) Source | Variance component | Error term | Expected Mean Square for Each Term (using unrestricted model) |
|-----------------|--------------------|------------|---|
| 1 Part | 8.683 | 3 | (4) + 2(3) + 4(1) |
| 2 Operator | 5.378 | 3 | (4) + 2(3) + 12(2) |
| 3 Part*Operator | -33.497 | 4 | (4) + 2(3) |
| 4 Error | 123.820 | | (4) |

Using MINITAB13>Stat>Quality Tools> Gauge R&R Study(Crossed), we obtain

| Source | VarComp | %Contribution (of VarComp) |
|-----------------|---------|-------------------------------|
| Total Gage R&R | 105.55 | 100.00 |
| Repeatability | 104.12 | 98.64 |
| Reproducibility | 1.44 | 1.36 |
| Operator | 1.44 | 1.36 |
| Part-To-Part | 0.00 | 0.00 |
| Total Variation | 105.55 | 100.00 |

| c) Source | StdDev (SD) | Study Var (5.15*SD) | %Study Var (%SV) |
|-----------------|----------------|------------------------|---------------------|
| Total Gage R&R | 10.2739 | 52.9106 | 100.00 |
| Repeatability | 10.2037 | 52.5491 | 99.32 |
| Reproducibility | 1.1989 | 6.1745 | 11.67 |
| Operator | 1.1989 | 6.1745 | 11.67 |
| Part-To-Part | 0.0000 | 0.0000 | 0.00 |
| Total Variation | 10.2739 | 52.9106 | 100.00 |

Overall most of the variability in the gauge comes from the measurement tool (repeatability).

8-40. a) Analysis of Variance for diameter

| Source | DF | SS | MS | F | P |
|---------------|----|---------|-------|------|-------|
| Part | 4 | 25.269 | 6.317 | 1.73 | 0.237 |
| Operator | 2 | 2.601 | 1.300 | 0.36 | 0.712 |
| Part*Operator | 8 | 29.290 | 3.661 | 0.61 | 0.759 |
| Error | 30 | 179.050 | 5.968 | | |
| Total | 44 | 236.209 | | | |

None of the terms is significant.

| b) Source | Variance component | Error term | Expected Mean Square for Each Term (using unrestricted model) |
|-----------------|--------------------|------------|---|
| 1 Part | 0.2951 | 3 | (4) + 3(3) + 9(1) |
| 2 Operator | -0.1574 | 3 | (4) + 3(3) + 15(2) |
| 3 Part*Operator | -0.7690 | 4 | (4) + 3(3) |
| 4 Error | 5.9683 | | (4) |

Using MINITAB13>Stat>Quality Tools> Gauge R&R Study(Crossed), we obtain

| Source | VarComp | %Contribution (of VarComp) |
|-----------------|---------|----------------------------|
| Total Gage R&R | 5.4826 | 98.34 |
| Repeatability | 5.4826 | 98.34 |
| Reproducibility | 0.0000 | 0.00 |
| Operator | 0.0000 | 0.00 |
| Part-To-Part | 0.0927 | 1.66 |
| Total Variation | 5.5754 | 100.00 |

| c) Source | StdDev (SD) | Study Var (5.15*SD) | %Study Var (%SV) |
|-----------------|-------------|---------------------|------------------|
| Total Gage R&R | 2.34150 | 12.0587 | 99.16 |
| Repeatability | 2.34150 | 12.0587 | 99.16 |
| Reproducibility | 0.00000 | 0.0000 | 0.00 |
| Operator | 0.00000 | 0.0000 | 0.00 |
| Part-To-Part | 0.30451 | 1.5682 | 12.90 |
| Total Variation | 2.36122 | 12.1603 | 100.00 |

Overall most of the variability in the gauge comes from the measurement tool (repeatability).

Supplementary Exercises

8-41. a)

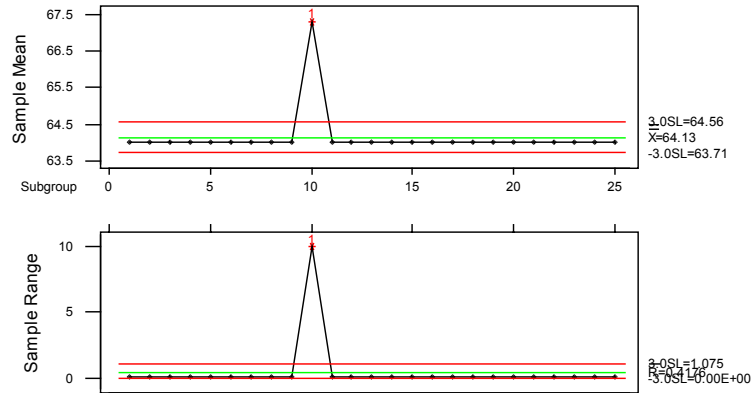
x chart \bar{R} chart

$$UCL = 64.56 \qquad UCL = 1.075$$

$$CL = 64.13 \qquad CL = 0.4176$$

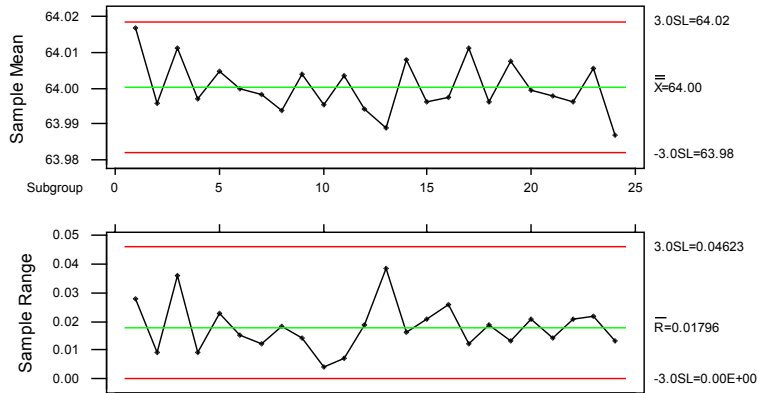
$$LCL = 63.71 \qquad LCL = 0$$

Xbar/R Chart for x1-x3



Observation # 10 is out of control.

Xbar/R Chart for x1-x3



The process is now in control.

$$b) \hat{\mu} = \bar{\bar{x}} = 64 \quad \hat{\sigma} = \frac{\bar{R}}{d_2} = \frac{0.01769}{1.693} = 0.0104$$

$$c) Cp = \frac{USL - LSL}{6\hat{\sigma}} = \frac{64.02 - 63.98}{6(0.0104)} = 0.641$$

The process does not meet the minimum capability level of $Cp \geq 1.33$.

d)

$$\begin{aligned} Cpk &= \min \left[\frac{USL - \bar{\bar{x}}}{3\hat{\sigma}}, \frac{\bar{\bar{x}} - LSL}{3\hat{\sigma}} \right] \\ &= \min \left[\frac{64.02 - 64}{3(0.0104)}, \frac{64 - 63.98}{3(0.0104)} \right] \\ &= \min [0.641, 0.641] \\ &= 0.641 \end{aligned}$$

e) In order to make this process a “six-sigma process”, the variance σ^2 would have to be decreased such that

$Cpk = 2.0$. The value of the variance is found by solving $Cpk = \frac{\bar{\bar{x}} - LSL}{3\sigma} = 2.0$ for σ :

$$\frac{64 - 61}{3\sigma} = 2.0$$

$$6\sigma = 64 - 61$$

$$\sigma = \frac{64 - 61}{6}$$

$$\sigma = 0.50$$

Therefore, the process variance would have to be decreased to $\sigma^2 = (0.50)^2 = 0.025$.

$$f) \hat{\sigma}_{\bar{x}} = \frac{\sigma}{\sqrt{n}} = \frac{0.0104}{\sqrt{3}} = 0.006$$

$$P(63.98 < \bar{X} < 64.02)$$

$$= P\left(\frac{63.98 - 64.005}{0.006} < \frac{\bar{X} - \mu}{\sigma_{\bar{x}}} < \frac{64.02 - 64.005}{0.006}\right)$$

$$= P(-4.17 < Z < 2.5)$$

$$= P(Z < 2.5) - P(Z < -4.17)$$

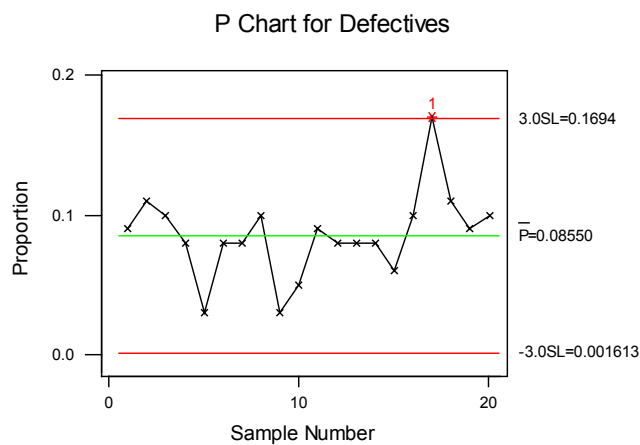
$$= 0.99379 - 0$$

$$= 0.99379$$

The probability that this shift will be detected on the next sample is $p = 1 - 0.99379 = 0.00621$

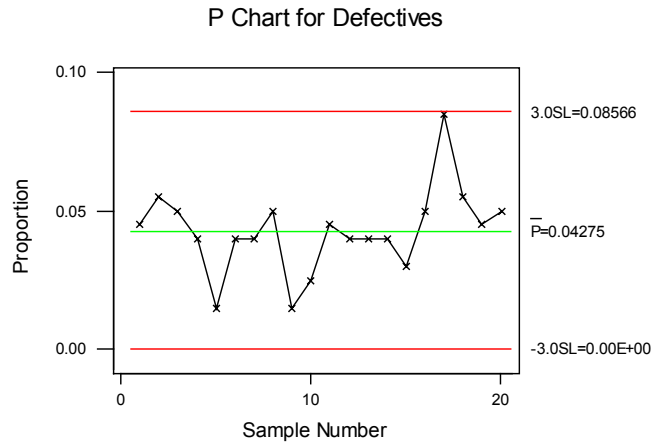
$$ARL = \frac{1}{p} = \frac{1}{0.00621} = 161$$

8-42. a)



The process is not in statistical control. Observation number 17 is beyond the UCL.

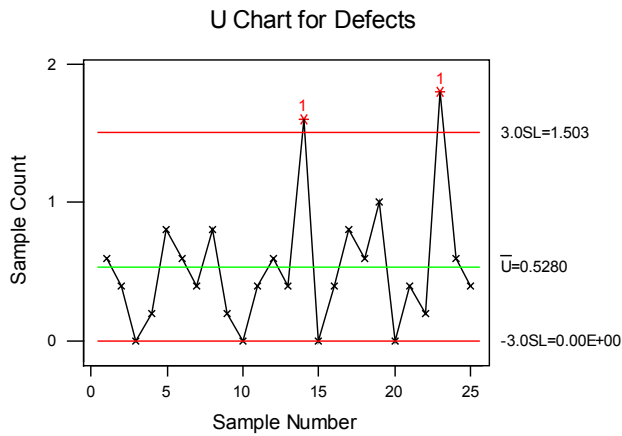
b)



The process appears to be in control.

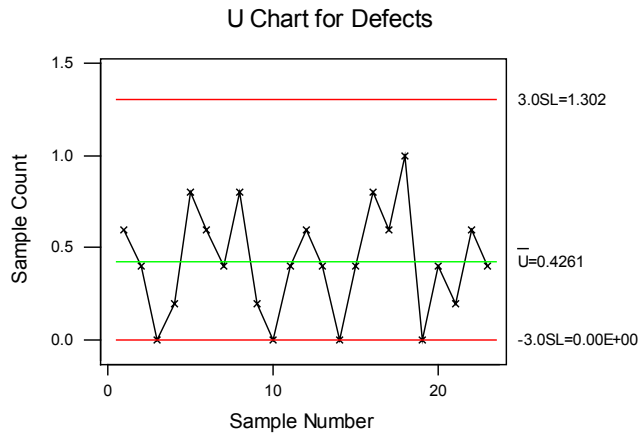
- c) A larger sample size with the same number of defective items will result in more narrow control limits. The control limits corresponding to the larger sample size are more sensitive.

8-43. a)



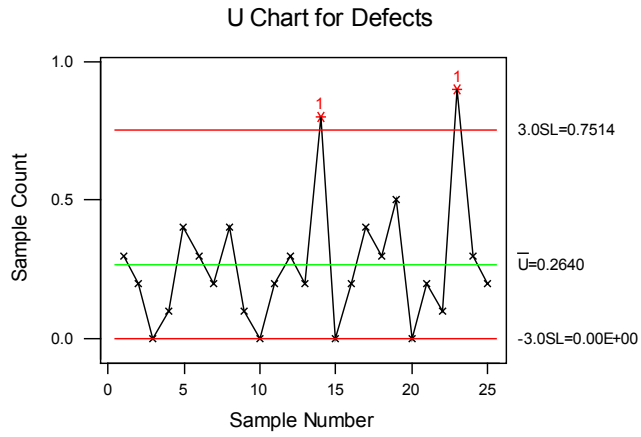
The process is out of control. Observation # 14 and 23 are beyond the control limits.

b)

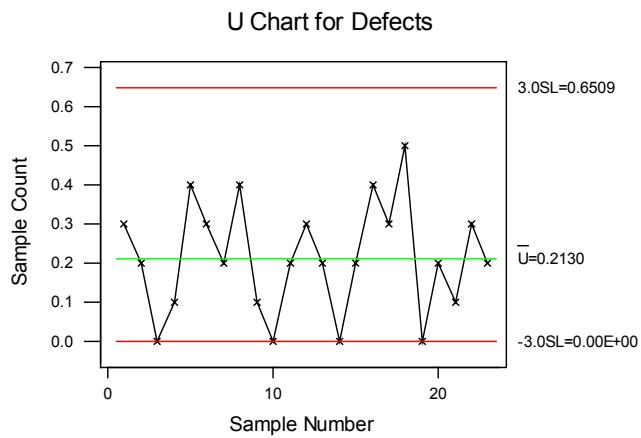


There are no points beyond the limits. The process is now in control.

c)

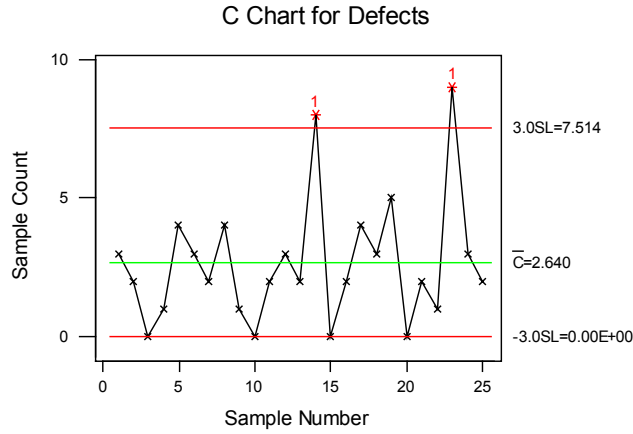


Revised U – Chart after removing point # 14 and 23.



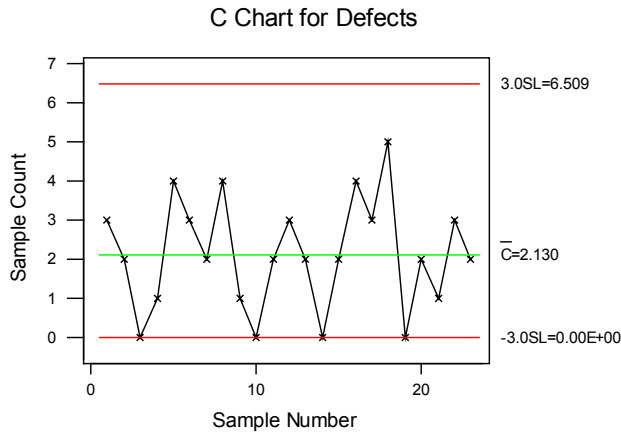
There are no points beyond the control limits. The process is now in control. Larger sample size narrows the control limits and causes more sample observations to be deemed out-of-control.

8-44. a)



There are points beyond the control limits. The process is out of control. The points are 14 and 23.

b)



There are no points beyond the control limits. The process is in control.

c) The control charts will not change since the sample size is not used in the c-chart.

8-45. a) Let p denote the probability that a point plots outside of the control limits when the mean has shifted from μ_0 to $\mu = \mu_0 + 1.75\sigma$. Then,

$$\begin{aligned}
& P(\text{LCL} < \bar{X} < \text{UCL}) \\
&= P\left(\mu_0 - \frac{3\sigma}{\sqrt{n}} < \bar{X} < \mu_0 + \frac{3\sigma}{\sqrt{n}}\right) \\
&= P\left(\frac{-1.75\sigma}{\sigma/\sqrt{4}} - 3 < \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < \frac{-1.75\sigma}{\sigma/\sqrt{4}} + 3\right) \\
&= P(-6.5 < Z < -0.5) \\
&= P(Z < -0.5) - P(Z < -6.5) \\
&= 0.30854 - 0 \\
&= 0.30854
\end{aligned}$$

Therefore, the probability the shift is undetected for three consecutive samples is $(1-p)^3 = (0.30854)^3 = 0.0294$.

b) If 2-sigma control limits were used, then

$$\begin{aligned}
1-p &= P(\text{LCL} < \bar{X} < \text{UCL}) \\
&= P\left(\mu_0 - \frac{2\sigma}{\sqrt{n}} < \bar{X} < \mu_0 + \frac{2\sigma}{\sqrt{n}}\right) \\
&= P\left(\frac{-1.75\sigma}{\sigma/\sqrt{n}} - 2 < \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < \frac{-1.75\sigma}{\sigma/\sqrt{n}} + 2\right) \\
&= P(-5.5 < Z < -1.5) \\
&= P(Z < -1.5) - P(Z < -5.5) \\
&= 0.06681
\end{aligned}$$

Therefore, the probability the shift is undetected for three consecutive samples is $(1-p)^3 = (0.06681)^3 = 0.0003$ 0.004.

c) The 2-sigma limits are more narrow than the 3-sigma limits. Since the 2-sigma limits have a smaller probability of a shift being undetected, it would be better than the 3-sigma limits for a mean shift of 1.5σ .

8-46. ARL = $1/p$ where p is the probability a point falls outside the control limits.

$$\begin{aligned}
& \text{a) } \mu = \mu_0 + \sigma \text{ and } n = 1 \\
& p = P(\bar{X} > \text{UCL}) + P(\bar{X} < \text{LCL}) \\
&= P\left(Z > \frac{\mu_0 + \frac{3\sigma}{\sqrt{n}} - \mu_0 - \sigma}{\sigma/\sqrt{n}}\right) + P\left(Z < \frac{\mu_0 - \frac{3\sigma}{\sqrt{n}} - \mu_0 - \sigma}{\sigma/\sqrt{n}}\right) \\
&= P(Z > 3 - \sqrt{n}) + P(Z < -3 - \sqrt{n}) \\
&= P(Z > 2) + P(Z < -4) \quad \text{when } n = 1 \\
&= 1 - P(Z < 2) + [1 - P(Z < 4)] \\
&= 1 - 0.97725 + [1 - 1] \\
&= 0.02275
\end{aligned}$$

Therefore, ARL = $1/p = 1/0.02275 = 43.9$.

b) $\mu = \mu_0 + 2\sigma$

$$\begin{aligned}
& P(\bar{X} > UCL) + P(\bar{X} < LCL) \\
&= P\left(Z > \frac{\mu_0 + \frac{3\sigma}{\sqrt{n}} - \mu_0 - 2\sigma}{\sigma/\sqrt{n}}\right) + P\left(Z < \frac{\mu_0 - \frac{3\sigma}{\sqrt{n}} - \mu_0 - 2\sigma}{\sigma/\sqrt{n}}\right) \\
&= P(Z > 3 - 2\sqrt{n}) + P(Z < -3 - 2\sqrt{n}) \\
&= P(Z > 1) + P(Z < -5) \quad \text{when } n = 1 \\
&= 1 - P(Z < 1) + [1 - P(Z < 5)] \\
&= 1 - 0.84134 + [1 - 1] \\
&= 0.15866
\end{aligned}$$

Therefore, $ARL = 1/p = 1/0.15866 = 6.30$.

c) $\mu = \mu_0 + 3\sigma$

$$\begin{aligned}
& P(\bar{X} > UCL) + P(\bar{X} < LCL) \\
&= P\left(Z > \frac{\mu_0 + \frac{3\sigma}{\sqrt{n}} - \mu_0 - 3\sigma}{\sigma/\sqrt{n}}\right) + P\left(Z < \frac{\mu_0 - \frac{3\sigma}{\sqrt{n}} - \mu_0 - 3\sigma}{\sigma/\sqrt{n}}\right) \\
&= P(Z > 3 - 3\sqrt{n}) + P(Z < -3 - 3\sqrt{n}) \\
&= P(Z > 0) + P(Z < -6) \quad \text{when } n = 1 \\
&= 1 - P(Z < 0) + [1 - P(Z < 6)] \\
&= 1 - 0.50 + [1 - 1] \\
&= 0.50
\end{aligned}$$

Therefore, $ARL = 1/p = 1/0.50 = 2.00$.

d) The ARL is decreasing as the magnitude of the shift increases from σ to 2σ to 3σ . The ARL will decrease as the magnitude of the shift increases since a larger shift is more likely to be detected earlier than a smaller shift.

8-47. a) Because $ARL = 370$, on the average we expect there to be one false alarm every 370 hours. Each 30-day month contains $30 \times 24 = 720$ hours of operation. Consequently, we expect $720/370 = 1.9$ false alarms each month.

b) The 2-sigma limits do reduce the ARL for detecting a shift in the mean of magnitude σ since the limits are narrower. The number of false alarms has increased using 2-sigma limits.

c) With 2-sigma limits the probability of a point plotting out of control is determined as follows, when

$$\mu = \mu_0 + \sigma$$

$$P(X > UCL) + P(X < LCL)$$

$$\begin{aligned}
&= P\left(\frac{X - \mu_0 - \sigma}{\sigma} > \frac{\mu_0 + 2\sigma - \mu_0 - \sigma}{\sigma}\right) + P\left(\frac{X - \mu_0 - \sigma}{\sigma} < \frac{\mu_0 - 2\sigma - \mu_0 - \sigma}{\sigma}\right) \\
&= P(Z > 1) + P(Z < -3) \\
&= 1 - P(Z < 1) + [1 - P(Z < 3)] \\
&= 1 - 0.84134 + 1 - 0.99865 \\
&= 0.160
\end{aligned}$$

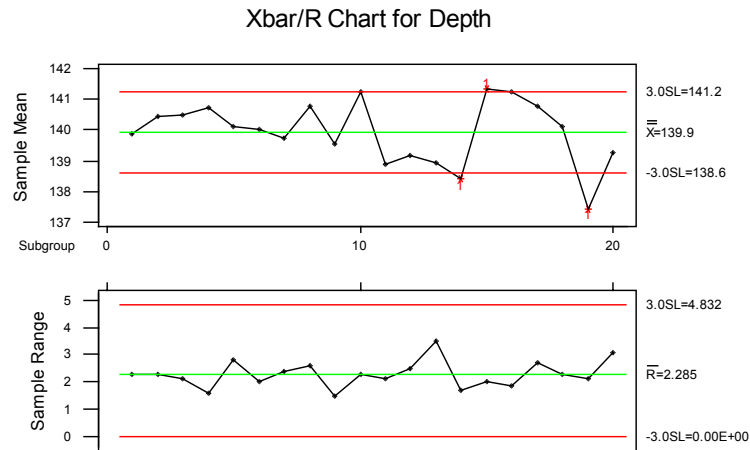
Therefore, $ARL = 1/p = 1/0.160 = 6.25$. The 2-sigma limits do reduce the ARL for detecting a shift in the mean of magnitude σ . The number of false alarms has increased using 2-sigma limits.

d) The in-control $ARL = 1/p$, where

$$\begin{aligned}
p &= P(X > UCL | \mu = \mu_0) + P(X < LCL | \mu = \mu_0) \\
&= P\left(\frac{X - \mu_0}{\sigma} > \frac{\mu_0 + 2\sigma - \mu_0}{\sigma}\right) + P\left(\frac{X - \mu_0}{\sigma} < \frac{\mu_0 - 2\sigma - \mu_0}{\sigma}\right) \\
&= P(Z > 2) + P(Z < -2) \\
&= 1 - P(Z < 2) + [1 - P(Z < 2)] \\
&= 2(1 - 0.97725) \\
&= 0.0455
\end{aligned}$$

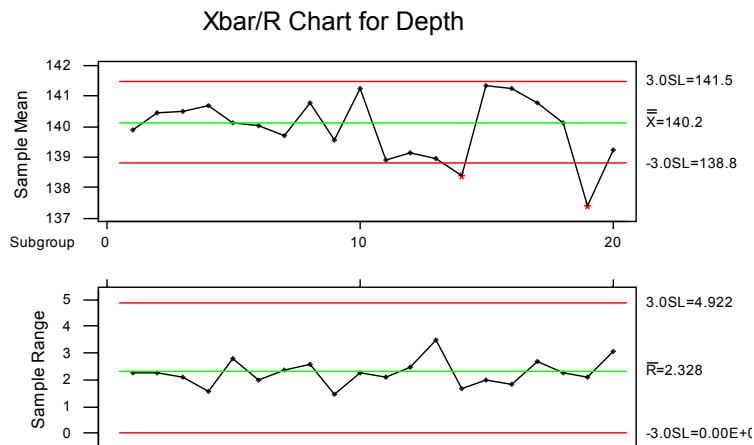
Therefore, $ARL = 1/0.0455 = 21.98$. The number of false alarms per month is $720/21.98 = 32.76$. This is an excessive number of false alarms (more than one per day) and 2-sigma limits are not recommended for routine production. Thus, this in-control ARL performance is probably not satisfactory.

8-48. a)



There are two points beyond the control limits. The process is out of control.

b)



The process is now in control.

The process standard deviation estimate is given by $\hat{\sigma} = \frac{\bar{R}}{d_2} = \frac{2.328}{2.326} = 1$

$$c) Cp = \frac{USL - LSL}{6\hat{\sigma}} = \frac{142 - 138}{6(1)} = 0.67$$

$$\begin{aligned} Cpk &= \min \left[\frac{USL - \bar{x}}{3\hat{\sigma}}, \frac{\bar{x} - LSL}{3\hat{\sigma}} \right] \\ &= \min \left[\frac{142 - 140.2}{3(1)}, \frac{140.2 - 138}{3(1)} \right] \\ &= \min[0.60, 0.73] \\ &= 0.60 \end{aligned}$$

Since Cp is less than unity, the natural tolerance limits lie outside the specification limits and many defective units will be produced.

Cp is slightly larger than Cpk indicating that the process is somewhat off center.

d) In order to make this process a “six-sigma process”, the variance σ^2 would have to be decreased such that

Cpk = 2.0. The value of the variance is found by solving $Cpk = \frac{\bar{x} - LSL}{3\sigma} = 2.0$ for σ :

$$\frac{140.2 - 138}{3\sigma} = 2.0$$

$$6\sigma = 140.2 - 138$$

$$\sigma = \frac{140.2 - 138}{6}$$

$$\sigma = 0.367$$

Therefore, the process variance would have to be decreased to $\sigma^2 = (0.367)^2 = 0.135$.

$$\begin{aligned} e) \hat{\sigma}_{\bar{x}} &= \frac{1}{\sqrt{5}} = 0.45 \\ p &= P(138.8 < \bar{x} < 141.5 | \mu = 139.7) \\ &= P\left(\frac{138.8 - 139.7}{0.45} < \frac{\bar{x} - \mu}{\sigma_{\bar{x}}} < \frac{141.5 - 139.7}{0.45}\right) \\ &= P(-2 < Z < 4) \\ &= P(Z < 4) - P(Z < -2) \\ &= 1 - 0.02275 \\ &= 0.97725 \end{aligned}$$

The probability that this shift will be detected on the next sample is $1 - p = 1 - 0.97725 = 0.02275$

$$ARL = \frac{1}{1 - p} = \frac{1}{0.02275} = 43.96$$

8-49. a) The $P(LCL < \hat{p} < UCL)$, when $p = 0.06$, is needed.

$$LCL = \bar{p} - 3\sqrt{\frac{\bar{p}(1 - \bar{p})}{n}} = 0.05 - 3\sqrt{\frac{0.05(1 - 0.05)}{100}} = -0.015 \rightarrow 0$$

$$UCL = \bar{p} + 3\sqrt{\frac{\bar{p}(1 - \bar{p})}{n}} = 0.05 + 3\sqrt{\frac{0.05(1 - 0.05)}{100}} = 0.115$$

Therefore, when $p = 0.06$

$$\begin{aligned}
 P(0 \leq \hat{P} \leq 0.115) &= P(\hat{P} \leq 0.115) \\
 &= P\left(\frac{\hat{P} - 0.06}{\sqrt{\frac{0.06(0.94)}{100}}} \leq \frac{0.115 - 0.06}{\sqrt{\frac{0.06(0.94)}{100}}}\right) \\
 &= P(Z \leq 2.32) \\
 &= 0.989830
 \end{aligned}$$

using the normal approximation to the distribution of \hat{P} . Therefore, the probability of detecting the shift on the first sample following the shift is $1 - 0.98983 = 0.01017$.

b) The probability that the control chart detects a shift to 0.06 on the second shift is $0.01017(0.98983) = 0.0101$.

c) $p = 0.08$

$$\begin{aligned}
 P(0 \leq \hat{P} \leq 0.115) &= P(\hat{P} \leq 0.115) \\
 &= P\left(\frac{\hat{P} - 0.10}{\sqrt{\frac{0.08(0.92)}{100}}} \leq \frac{0.115 - 0.08}{\sqrt{\frac{0.08(0.92)}{100}}}\right) \\
 &= P(Z \leq 1.29) \\
 &= 0.901475
 \end{aligned}$$

using the normal approximation to the distribution of \hat{P} . Therefore, the probability of detecting the shift on the first sample following the shift is $1 - 0.901475 = 0.098525$.

The probability that the control chart detects a shift to 0.10 on the second shift is $0.901475(0.098525) = 0.0888$.

d) A larger shift is generally easier to detect. Therefore, we should expect a shift to 0.08 to be detected quicker than a shift to 0.06.

8-50. $\bar{u} = 8$

a) $n = 5$

$$UCL = \bar{u} + 3\sqrt{\frac{\bar{u}}{n}} = 8 + 3\sqrt{\frac{8}{5}} = 11.79$$

$$LCL = \bar{u} - 3\sqrt{\frac{\bar{u}}{n}} = 8 - 3\sqrt{\frac{8}{5}} = 4.21$$

$$\begin{aligned}
 P(\bar{U} > 11.79 \text{ when } \lambda = 16) &= P\left(Z > \frac{11.79 - 16}{\sqrt{16/5}}\right) \\
 &= P(Z > -2.35) \\
 &= 1 - P(Z < -2.35) \\
 &= 1 - 0.009387 \\
 &= 0.990613
 \end{aligned}$$

$$\begin{aligned}
 P(\bar{U} < 4.21 \text{ when } \lambda = 16) &= P\left(Z < \frac{4.21 - 16}{\sqrt{16/5}}\right) \\
 &= P(Z < -6.59) \\
 &= 0
 \end{aligned}$$

The probability is then 0.990613
 b) $n = 8$

$$UCL = \bar{u} + 3\sqrt{\frac{\bar{u}}{n}} = 8 + 3\sqrt{\frac{8}{8}} = 11$$

$$LCL = \bar{u} - 3\sqrt{\frac{\bar{u}}{n}} = 8 - 3\sqrt{\frac{8}{8}} = 5$$

$$\begin{aligned} P(\bar{U} > 11 \text{ when } \lambda = 16) &= P\left(Z > \frac{11-16}{\sqrt{\frac{16}{8}}}\right) \\ &= P(Z > -3.54) \\ &= 0.9998 \end{aligned}$$

$$\begin{aligned} P(\bar{U} < 5 \text{ when } \lambda = 16) &= P\left(Z < \frac{5-16}{\sqrt{\frac{16}{8}}}\right) \\ &= P(Z < -7.78) \\ &= 0 \end{aligned}$$

The probability is then 0.9998

8-51. $\bar{u} = 10$
 a) $n = 3$

$$UCL = \bar{u} + 3\sqrt{\frac{\bar{u}}{n}} = 10 + 3\sqrt{\frac{10}{3}} = 15.48$$

$$LCL = \bar{u} - 3\sqrt{\frac{\bar{u}}{n}} = 10 - 3\sqrt{\frac{10}{3}} = 4.52$$

$$\begin{aligned} P(\bar{U} > 15.48 \text{ when } \lambda = 14) &= P\left(Z > \frac{15.48-14}{\sqrt{\frac{14}{3}}}\right) \\ &= P(Z > 0.685) \\ &= 1 - P(Z < 0.685) \\ &= 1 - 0.7533 \\ &= 0.2467 \end{aligned}$$

and

$$P(\bar{U} < 4.52) = P\left(Z < \frac{4.52-14}{\sqrt{\frac{14}{3}}}\right) = 0$$

The probability is then 0.2467

b) $n = 6$

$$UCL = \bar{u} + 3\sqrt{\frac{\bar{u}}{n}} = 10 + 3\sqrt{\frac{10}{6}} = 13.87$$

$$LCL = \bar{u} - 3\sqrt{\frac{\bar{u}}{n}} = 10 - 3\sqrt{\frac{10}{6}} = 6.13$$

$$\begin{aligned} P(\bar{U} > 13.87 \text{ when } \lambda = 14) &= P\left(Z > \frac{13.87-14}{\sqrt{\frac{14}{6}}}\right) \\ &= P(Z > -0.085) \\ &= 1 - 0.4661 \end{aligned}$$

$$\begin{aligned}
 &= 0.5339 \\
 P(\bar{U} < 6.13 \text{ when } \lambda = 14) &= P\left(Z < \frac{6.13 - 14}{\sqrt{\frac{14}{6}}}\right) \\
 &= P(Z < -5.15) \\
 &= 0
 \end{aligned}$$

The probability is then 0.5339

8-52.

$$CL = \mu$$

$$UCL = \mu + 2\frac{\sigma}{\sqrt{n}}$$

$$LCL = \mu - 2\frac{\sigma}{\sqrt{n}}$$

$$P\left(\bar{X} > \mu + 2\frac{\sigma}{\sqrt{n}}\right) = P\left(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} > 2\right)$$

$$= P(Z > 2)$$

$$= 1 - P(Z < 2)$$

$$= 1 - 0.97725$$

$$= 0.02275$$

and

$$P\left(\bar{X} < \mu - 2\frac{\sigma}{\sqrt{n}}\right) = P\left(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < -2\right)$$

$$= P(Z < -2)$$

$$= 1 - P(Z < 2)$$

$$= 1 - 0.97725$$

$$= 0.02275$$

The answer is $0.02275 + 0.02275 = 0.0455$.

The answer for 3-sigma control limits is 0.0027. The 3-sigma control limits result in much fewer false alarms.

8-53.

$$CL = \mu$$

$$UCL = \mu + k\frac{\sigma}{\sqrt{n}}$$

$$LCL = \mu - k\frac{\sigma}{\sqrt{n}}$$

$$\begin{aligned}
P\left(\bar{X} > \mu + k \frac{\sigma}{\sqrt{n}}\right) &= P\left(\frac{\bar{X} - \mu - \delta}{\sigma/\sqrt{n}} > k - \frac{\delta}{\sigma/\sqrt{n}}\right) \\
&= P\left(Z > k - \frac{\delta}{\sigma/\sqrt{n}}\right) \\
&= 1 - \Phi\left(k - \frac{\delta}{\sigma/\sqrt{n}}\right)
\end{aligned}$$

$$\begin{aligned}
P\left(\bar{X} < \mu + k \frac{\sigma}{\sqrt{n}}\right) &= P\left(\frac{\bar{X} - \mu - \delta}{\sigma/\sqrt{n}} < -k - \frac{\delta}{\sigma/\sqrt{n}}\right) \\
&= P\left(Z < -k - \frac{\delta}{\sigma/\sqrt{n}}\right) \\
&= \Phi\left(-k - \frac{\delta}{\sigma/\sqrt{n}}\right)
\end{aligned}$$

The answer is

$$1 - \Phi\left(k - \frac{\delta\sqrt{n}}{\sigma}\right) + \Phi\left(-k - \frac{\delta\sqrt{n}}{\sigma}\right).$$

8-54. From Exercise 8-52, $p = 0.0455$.

a) The probability of a false alarm is the probability that \bar{X} is outside the control limits when there is no shift, i.e., 0.0455. The probability there is not a false alarm on the first sample is $1 - 0.0455 = 0.9545$. Hence, the probability of a false alarm on the second sample but not on the first (assuming the two samples are independent) is $(0.9545)(0.0455) = 0.04343$.

b) Assuming all the samples are independent, the probability that there is not a false alarm in the first three samples is $(0.9545)^3 = 0.86962$.

8-55. $C_p = 2$ but $\mu = USL + 3\sigma$

$$P(X < USL) = P\left(Z < \frac{(\mu - 3\sigma) - \mu}{\sigma}\right) = P(Z < -3) = 0.00135$$

8-56. a) From Prob. 8-37, we have $\hat{\sigma}_{gauge} = 0.2872$. Thus the P/T ratio is

$$P/T \text{ ratio} = \frac{6\hat{\sigma}_{gauge}}{USL - LSL} = \frac{6(0.2872)}{2.8 - 1.2} = 1.077$$

b) Because P/T ratio is 1.077 higher than 0.1, the gauge is not adequate.

8-57. a) From Prob. 8-38, we have $\hat{\sigma}_{gauge} = 4.7588$. Thus the P/T ratio is

$$P/T \text{ ratio} = \frac{6\hat{\sigma}_{gauge}}{USL - LSL} = \frac{6(4.7588)}{665 - 635} = 0.9518$$

b) Because P/T ratio is 0.9518 higher than 0.1, the gauge is not adequate.