

Chapter 3

	y	$p_Y(y)$
	0	0.6
	20	0.3
	80	0.1
	ow	0

$$(b) \quad E(Y) = \sum_y y \cdot p_Y(y) = 14$$

$$V(Y) = \sum_y y^2 \cdot p_Y(y) - (14)^2 = 564$$

3-2. Let Profit = $P = 10 + 2X$

$$P(P \leq p) = P(10 + 2X \leq p) = P\left(X \leq \frac{p-10}{2}\right) = \int_0^{\frac{p-10}{2}} \frac{x}{18} dx$$

$$F_P(p) = \frac{x^2}{36} \Big|_0^{\frac{p-10}{2}} = \frac{1}{144}(p^2 - 20p + 100)$$

$$f_P(p) = \frac{1}{72}(p-10); 10 \leq p \leq 22$$

$$= 0; \text{ow}$$

3-3. (a) $P(T < 1) = 1 - e^{-1/4} = 0.221$

(b) $E[P] = 200 - 200P(T < 1) = 155.80$

	x	$p_X(x)$	$y = 2000(12 - x)$	$p_Y(y)$
	10	0.1	4000	0.1
	11	0.3	2000	0.3
	12	0.4	0	0.4
	13	0.1	-2000	0.1
	14	0.1	-4000	0.1
	ow	0	ow	0

(b) $E(X) = 10(0.1) + 11(0.3) + 12(0.4) + 13(0.1) + 14(0.1) = 11.8 \text{ days}$

$$V(X) = 10^2(0.1) + 11^2(0.3) + 12^2(0.4) + 13^2(0.1) + 14^2(0.1) - (11.8)^2$$

$$= 1.16 \text{ days}^2$$

$$\begin{aligned} E(Y) &= (4000)(0.1) + (2000)(0.3) + 0(0.4) + (-2000)(0.1) \\ &\quad + (-4000)(0.1) = \$400 \end{aligned}$$

$$\begin{aligned} V(Y) &= (4000)^2(0.1) + (2000)^2(0.3) + 0^2(0.4) + (-2000)^2(0.1) \\ &\quad + (-4000)^2(0.1) - 400^2 = 4,640,000(\$^2) \end{aligned}$$

$$\begin{aligned} 3-5. \quad F_Z(z) &= P(Z \leq z) = P(X^2 \leq z) = P(|X| \leq \sqrt{z}) \\ &= P(0 \leq X \leq \sqrt{z}) = \int_0^{\sqrt{z}} 2xe^{-x^2} dx \end{aligned}$$

Let $u = x^2$, $du = 2x dx$, so

$$F_Z(z) = \int_0^z e^{-u} du = 1 - e^z$$

$$f_Z(z) = e^{-z}; z \geq 0$$

$$= 0; \text{ow}$$

$$3-6. \quad (\text{a}) \quad E(D_i) = \sum_{d=0}^9 d \cdot p_{D_i}(d) = \frac{1}{10}(1 + 2 + 3 + 4 + 5 + 6 + 7 + 8 + 9) = 4.5$$

$$\begin{aligned} (\text{b}) \quad V(D_i) &= \sum_{d=0}^9 d^2 \cdot p_{D_i}(d) - (4.5)^2 \\ &= \frac{1}{10}[1^2 + 2^2 + 3^2 + 4^2 + 5^2 + 6^2 + 7^2 + 8^2 + 9^2] - (20.25) = 8.25 \end{aligned}$$

(c)	d	y	y	$p_Y(y)$
	0	4	0	0.2
	1	3	1	0.2
	2	2	2	0.2
	3	1	3	0.2
	4	0	4	0.2
	5	0	ow	0
	6	1		
	7	2		
	8	3		
	9	4		

$$E(Y) = \frac{2}{10}(1 + 2 + 3 + 4) = 2$$

$$V(Y) = \frac{2}{10}(1^2 + 2^2 + 3^2 + 4^2) - 4 = 2$$

$$3-7. \quad R = \text{Revenue/Gal}$$

$$R = 0.92; A < 0.7$$

$$= 0.98; A \geq 0.7$$

$$\begin{aligned} E(R) &= 0.92 P(A < 0.7) + 0.98 P(A \geq 0.7) \\ &= 0.92(0.7) + 0.98(0.3) = 93.8\text{¢/gal.} \end{aligned}$$

$$3-8. \quad M_X(t) = \int_{\beta}^{\infty} e^{tx} \frac{1}{\theta} e^{-\frac{1}{\theta}(x-\beta)} dx = \frac{1}{\theta} e^{\beta/\theta} \int_{\beta}^{\infty} e^{-x(\frac{1}{\theta}-t)} dx$$

$$= \frac{1}{\theta} \left(\frac{1}{\theta} - t \right)^{-1} e^{\beta t}, \text{ for } \frac{1}{\theta} - t > 0$$

$$M'_X(t) = \frac{1}{\theta} \left(\frac{1-\theta t}{\theta} \right)^{-2} e^{\beta t} \left[\frac{1-\theta t}{\theta} \beta + 1 \right]$$

$$M''_X(t) = \frac{1}{\theta} \left(\frac{1-\theta t}{\theta} \right)^{-3} e^{\beta t} \left[\left(\frac{1-\theta t}{\theta} \right)^2 \beta^2 + \left(\frac{1-\theta t}{\theta} \right) \beta + 2 + \left(\frac{1-\theta t}{\theta} \right) \beta \right]$$

$$E(X) = M'_X(0) = \beta + \theta$$

$$V(X) = M''_X(0) - (\beta + \theta)^2 = \theta^2$$

$$3-9. \quad (a) \quad F_Y(y) = P(2X^2 \leq y) = P \left[-\sqrt{\frac{y}{2}} \leq X \leq +\sqrt{\frac{y}{2}} \right]$$

$$= P \left(0 \leq X \leq \sqrt{\frac{y}{2}} \right)$$

$$= \int_0^{\sqrt{\frac{y}{2}}} e^{-x} dx = -e^{-x} \Big|_0^{\sqrt{\frac{y}{2}}} = 1 - e^{-\sqrt{\frac{y}{2}}}$$

$$f_Y(y) = F'_Y(y) = \frac{1}{4} \left(\frac{y}{2} \right)^{-1/2} \cdot e^{-(y/2)^{1/2}}; y > 0$$

$$= 0; \text{ otherwise}$$

$$(b) \quad F_V(v) = P(X^{1/2} \leq v) = P(X \leq v^2)$$

$$= \int_0^{v^2} e^{-x} dx = -e^{-x}|_0^{v^2} = 1 - e^{-v^2}$$

$$f_V(v) = F'_V(v) = 2ve^{-v^2}; v > 0$$

$$= 0; \text{otherwise}$$

$$(c) \quad F_U(u) = P(\ln(X) \leq u) = P(X \leq e^u) = \int_0^{e^u} e^{-x} dx$$

$$= 1 - e^{-e^u}$$

$$f_U(u) = F'_U(u) = e^{-(e^u-u)}; u > 0$$

$$= 0; \text{otherwise}$$

3–10. Note that as stated $Y > y_0 \Rightarrow$ signal read (not $Y > |y_0|$), so

$$P_Y(Y > y) = \int_{\tan^{-1}y}^{\pi/2} \frac{1}{2\pi} dx + \int_{\tan^{-1}y+\pi}^{3\pi/2} \frac{1}{2\pi} dx$$

$$F_Y(y) = 1 - P_Y(Y > y)$$

$$= \frac{1}{2} + \frac{1}{\pi} \tan^{-1}y; y \geq 0$$

$$f_Y(y) = \frac{1}{\pi} \left(\frac{1}{1+y^2} \right); -\infty < y < +\infty$$

Note the symmetry in $(\frac{1}{1+y^2})$.

3–11. Let

$S = \text{Stock Level}$

$$L(X, S) = 0.5(X - S), X > S$$

$$= 0.25(S - X), X \leq S$$

$$E(L(X, S)) = \int_{10^6}^S 0.25(S - X) \cdot 10^6 dx + \int_S^{2 \cdot 10^6} 0.5(X - S) \cdot 10^{-6} dx$$

$$\frac{dE[L(X, S)]}{dS} = \frac{6}{8} 10^{-6} S - \frac{5}{4} = 0 \Rightarrow S = \frac{5}{3} 10^6$$

$$3-12. \quad \mu_G = E(G) = \int_0^4 g f_G(g) dg = \int_0^4 \frac{g^2}{8} dg = 8/3.$$

$$E(G^2) = \int_0^4 \frac{g^3}{8} dg = 8.$$

$$\sigma_G^2 = V(G) = E(G^2) - (E(G))^2 = 8/9.$$

$$H(G) = (3 + 0.05G)^2$$

$$H(\mu_G) = (3 + 0.05\mu_G)^2$$

$$H'(\mu_G) = (0.1)(3 + 0.05\mu_G)$$

$$H''(\mu_G) = 0.005.$$

Thus,

$$\mu_A \approx H(\mu_G) + \frac{1}{2}H''(\mu_G)\sigma_G^2 = H(8/3) + \frac{1}{2}(0.005)\frac{8}{9} = 9.82$$

and

$$\sigma_A^2 \approx (H'(\mu_G))^2\sigma_G^2 = 0.082.$$

$$3-13. \quad (a) \quad F_Y(y) = P(4 - X^2 \leq y) = P(X^2 \geq 4 - y) \\ = P(X \leq -\sqrt{4-y}) + P(X \geq \sqrt{4-y}) \\ = 0 + \int_{\sqrt{4-y}}^2 dx = 2 - \sqrt{4-y}$$

$$f_Y(y) = \frac{1}{2}(4-y)^{-1/2}; 0 \leq y \leq 3$$

= 0; otherwise

$$(b) \quad F_Y(y) = P(e^x \leq y) = P(X \leq \ln(y))$$

$$= \int_1^{\ln(y)} dx = \ln(y) - 1$$

$$f_Y(y) = \frac{1}{y}; e^1 \leq y \leq e^2$$

= 0; otherwise

3–14. $y = \frac{3}{(1+x)^2}$

$$x = \sqrt{3}y^{-1/2} - 1$$

$$\frac{dx}{dy} = -\frac{\sqrt{3}}{2}y^{-3/2}$$

$$\left| \frac{dx}{dy} \right| = \frac{\sqrt{3}}{2}y^{-3/2}$$

$$f_Y(y) = f_X(x) \left| \frac{dx}{dy} \right|$$

$$= f_X(\sqrt{3}y^{-1/2} - 1) \frac{\sqrt{3}}{2}y^{-3/2}$$

$$= \exp[-\sqrt{3}y^{-1/2} + 1] \frac{\sqrt{3}}{2}y^{-3/2}; \quad 0 \leq y \leq 3$$

3–15. With equal probability $p_X(1) = \dots = p_X(6) = \frac{1}{6}$,

$$M_X(t) = \sum_{x=1}^6 \left(\frac{1}{6} \right) e^{tx}$$

$$E(X) = M'_X(0) = \frac{7}{2}$$

$$V(X) = M''_X(0) - [M'_X(0)]^2 = \frac{35}{12}$$

3–16. (a) Using the substitution $y = bx^2$, we obtain

$$\begin{aligned} 1 &= \int_0^\infty ax^2 e^{-bx^2} dx \\ &= \int_0^\infty a \frac{y}{b} e^{-y} \frac{dy}{2\sqrt{by}} \\ &= \frac{a}{2b^{3/2}} \int_0^\infty y^{(3/2)-1} e^{-y} dy \\ &= \frac{a}{2b^{3/2}} \Gamma\left(\frac{3}{2}\right) = \frac{a}{2b^{3/2}} \frac{1}{2} \Gamma\left(\frac{1}{2}\right) = \frac{a\sqrt{\pi}}{4b^{3/2}} \end{aligned}$$

Thus,

$$a = \frac{4b^{3/2}}{\sqrt{\pi}}.$$

(b) Similarly,

$$\mu_X = E(X) = \int_0^\infty ax^3 e^{-bx^2} dx = \frac{a}{2b^2} \Gamma(2) = \frac{2}{\sqrt{\pi b}}$$

and

$$E(X^2) = \frac{a}{2b^{5/2}} \Gamma\left(\frac{5}{2}\right) = \frac{3}{2b}.$$

These facts imply that

$$\sigma_X^2 = E(X^2) - [E(X)]^2 = \frac{3}{2b} - \frac{4}{\pi b} = \frac{3\pi - 8}{2\pi b}.$$

Now we have

$$\begin{aligned} H(x) &= 18x^2 \\ H(\mu_X) &= 18\mu_X^2 = 18 \cdot \frac{4}{\pi b} = \frac{72}{\pi b} \\ H'(\mu_X) &= 36\mu_X = 36 \cdot \frac{2}{\sqrt{\pi b}} = \frac{72}{\sqrt{\pi b}} \\ H''(\mu_X) &= 36 \end{aligned}$$

Then

$$\begin{aligned} \mu_Y &\doteq H(\mu_X) + \frac{1}{2}H''(\mu_X) \cdot \sigma_X^2 \\ &= \frac{72}{\pi b} + \frac{1}{2} \cdot 36 \cdot \frac{3\pi - 8}{2\pi b} = \frac{27}{b} \end{aligned}$$

and

$$\begin{aligned} \sigma_Y^2 &\doteq (H'(\mu_X))^2 \sigma_X^2 \\ &= \left(\frac{72}{\sqrt{\pi b}}\right)^2 \left(\frac{3\pi - 8}{2\pi b}\right) = \frac{2592(3\pi - 8)}{\pi^2 b^2} \end{aligned}$$

$$3-17. \quad E(Y) = 1, V(Y) = 1 \quad H(Y) = \sqrt{Y^2 + 36}$$

$$E(X) \cong H(\mu_Y) + \frac{1}{2}H''(\mu_Y)\sigma_Y^2$$

$$V(X) \cong [H'(\mu_Y)]^2 \sigma_Y^2$$

$$H'(y) = y(y^2 + 36)^{-1/2}, \quad H''(y) = 36(y^2 + 36)^{-3/2}$$

$$E(X) \cong (37)^{1/2} + \frac{1}{2}(36)(37)^{-3/2} \doteq 6.16$$

$$V(X) \cong (37)^{-1} = 0.027027$$

$$3-18. E(P) = \int_0^1 (1+3r)6r(1-r) dr = \int_0^1 (6r + 12r^2 - 18r^3) dr = \frac{5}{2}$$

Since $H(r) = 1 + 3r$ is increasing,

$$\begin{aligned} f_P(p) &= 6 \left(\frac{p-1}{3} \right) \left(1 - \frac{p-1}{3} \right) \left(\frac{1}{3} \right) = \frac{2}{9}(5p - p^2 - 4); 1 < p < 4 \\ &= 0; \text{ otherwise} \end{aligned}$$

$$\begin{aligned} 3-19. M_X(t) &= \int_0^\infty e^{tx} \cdot 4xe^{-2x} dx \\ &= \int_0^\infty 4xe^{x(t-2)} dx = \left(1 - \frac{t}{2} \right)^{-2} \end{aligned}$$

$$E(X) = M'_X(0) = 1$$

$$V(X) = M''_X(0) - [M'_X(0)]^2 = \frac{1}{2}$$

$$\begin{aligned} 3-20. \quad V &= \frac{\pi \cdot X^2}{4} \cdot 1 \\ E(V) &= \frac{\pi}{4} E(X^2) = \frac{\pi}{4} [V(X) + (E(X))^2] \\ &= \frac{\pi}{4} \cdot [25 \cdot 10^{-6} + 2^2] = 3.14162 \end{aligned}$$

$$\begin{aligned} 3-21. M_Y(t) &= E(e^{tY}) = E(e^{t(aX+b)}) \\ &= e^{tb} E(e^{(at)X}) \\ &= e^{tb} M_X(at) \end{aligned}$$

$$3-22. \quad (\text{a}) \quad \int_0^1 k(1-x)^{a-1} x^{b-1} dx = 1.$$

Let $\Gamma(p) = \int_0^\infty u^{p-1} e^{-u} du$ define the gamma function. Integrating by parts and applying L'Hôpital's Rule, it can be shown that $\Gamma(p+1) = p \cdot \Gamma(p)$.

Make a variable change $u = v^2$ so that $\Gamma(\frac{1}{2}) = \sqrt{\pi}$.

Working with $\Gamma(p)$, let $u = v^2$ so $du = 2v \, dv$ and

$$\Gamma(p) = 2 \int_0^\infty (v^2)^{p-1} e^{-v^2} v \, dv = 2 \int_0^\infty v^{2p-1} e^{-v^2} \, dv.$$

$$\begin{aligned} \text{Then } \Gamma(a) \cdot \Gamma(b) &= 4 \int_0^\infty s^{2a-1} e^{-s^2} \, ds \int_0^\infty t^{2b-1} e^{-t^2} \, dt \\ &= 4 \int_0^\infty \int_0^\infty s^{2a-1} t^{2b-1} e^{-(s^2+t^2)} \, ds \, dt \end{aligned}$$

Let $s = \rho \cos \theta$, $t = \rho \sin \theta$, so the Jacobian = ρ

$$\begin{aligned} \Gamma(a) \cdot \Gamma(b) &= 4 \int_0^\infty \int_0^{\pi/2} (\rho \cos \theta)^{2a-1} (\rho \sin \theta)^{2b-1} e^{-\rho^2} \rho \, d\theta \, d\rho \\ &= 4 \int_0^\infty \rho^{2a+2b-1} e^{-\rho^2} \, d\rho \int_0^{\pi/2} (\cos \theta)^{2a-1} (\sin \theta)^{2b-1} \, d\theta \end{aligned}$$

Substitute $\rho^2 = y$ in the first integral and $\sin^2 \theta = x$ in the second.

$$\begin{aligned} \Gamma(a)\Gamma(b) &= \int_0^\infty y^{a+b-1} e^{-y} \, dy \int_0^1 x^{b-1} (1-x)^{a-1} \, dx \\ &= \Gamma(a+b) \int_0^1 x^{b-1} (1-x)^{a-1} \, dx \\ \text{so } \int_0^1 x^{b-1} (1-x)^{a-1} \, dx &= \frac{\Gamma(a) \cdot \Gamma(b)}{\Gamma(a+b)} \Rightarrow k = \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \end{aligned}$$

$$\begin{aligned}
(b) \quad E(X^k) &= \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \int_0^1 x^{b+k-1} (1-x)^{a-1} \\
&= \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \int_0^1 x^{(b+k)-1} (1-x)^{a-1} dx \\
&= \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \cdot \frac{\Gamma(a)\Gamma(b+k)}{\Gamma(a+b+k)}.
\end{aligned}$$

Since $\Gamma(1) = 1$ we obtain $\Gamma(p+1) = p!$ for p a positive integer.

$$\begin{aligned}
\text{Then } E(X) &= \frac{(a+b-1)!}{(a-1)!(b-1)!} = \frac{b}{a+b} \\
E(X^2) &= \frac{[(a+b)-1]!}{(a-1)!(b-1)!} \cdot \frac{(a-1)!(b+1)!}{(a+b+1)!} = \frac{b(b+1)}{(a+b+1)(a+b)} \\
\text{so } V(X) &= \frac{ab}{(a+b)^2(a+b+1)}
\end{aligned}$$

$$3-23. \quad M_X(t) = \frac{1}{2} + \frac{1}{4}e^t + \frac{1}{8}e^{2t} + \frac{1}{8}e^{3t}$$

$$M'_X(t) = \frac{1}{4}e^t + \frac{1}{4}e^{2t} + \frac{3}{8}e^{3t}$$

$$M''_X(t) = \frac{1}{4}e^t + \frac{1}{2}e^{2t} + \frac{9}{8}e^{3t}$$

$$(a) \quad E(X) = M'_X(0) = \frac{7}{8}$$

$$V(Y) = M''_X(0) - [M'_X(0)]^2 = \frac{15}{8} - \left(\frac{7}{8}\right)^2 = \frac{71}{64}$$

(b)	x	y	y	$p_Y(y)$
	0	4	0	$\frac{1}{8}$
	1	1	1	$\frac{3}{8}$
	2	0	4	$\frac{1}{2}$
	3	1	ow	0
	ow	0		

$$F_Y(y) = 0; y < 0$$

$$\begin{aligned} &= \frac{1}{8}; 0 \leq y < 1 \\ &= \frac{1}{2}; 1 \leq y < 4 \\ &= 1; y \geq 4 \end{aligned}$$

$$3-24. \mu_3 = E(X - \mu'_1)^3$$

$$\begin{aligned} &= E(X^3) - 3\mu'_1 \cdot E(X^2) + 3\mu'_1 \cdot E(X) - (\mu'_1)^3 \\ &= \mu'_3 - 3\mu'_1 \mu'_2 + 2(\mu'_1)^3. \end{aligned}$$

A r.v. is symmetric about a if $f(x+a) = f(-x+a)$ and we assume X continuous. Since $X - a$ and $a - X$ have the same p.d.f., $E(X - a) = E(a - X)$ or $E(X) - a = a - E(X)$, so $E(X) = a = \mu'_1$. Now,

$$\begin{aligned} E(X - \mu'_1)^r &= E(X - a)^r = \int_{-\infty}^{\infty} (x - a)^r f(x) dx \\ &= \int_{-\infty}^a (x - a)^r f(x) dx + \int_0^{\infty} (x - a)^r f(x) dx \\ &= \int_{-\infty}^0 y^r f(y + a) dy + \int_0^{\infty} y^r f(y + a) dy \\ &= - \int_{-\infty}^0 (1 - y)^r f(-y + a) dy + \int_0^{\infty} y^r f(y + a) dy \\ &= \int_0^{\infty} (-y)^r f(y + a) dy + \int_0^{\infty} y^r f(y + a) dy \\ &= (-1)^r \int_0^{\infty} y^r f(y + a) dy + \int_0^{\infty} y^r f(y + a) dy \\ &= 0 \text{ for odd } r. \end{aligned}$$

Thus $E(X - \mu'_1)^3 = 0$.

3–25. If $\int_{-\infty}^{\infty} |x|^r f(x) dx = k < \infty$, then $\int_{-\infty}^{\infty} |x|^n f(x) dx < \infty$, where $0 \leq n < r$.

Proof: If $|x| \leq 1$, then $|x|^n \leq 1$ and if $|x| > 1$, then $|x|^n \leq |x|^r$.

Then

$$\begin{aligned}\int_{-\infty}^{\infty} |x|^n f(x) dx &= \int_{|x| \leq 1} |x|^n \cdot f(x) dx + \int_{|x| > 1} |x|^n \cdot f(x) dx \\ &\leq \int_{|x| \leq 1} 1 f(x) dx + \int_{|x| > 1} |x|^r f(x) dx \leq 1 + k < \infty\end{aligned}$$

3–26. $M_X(t) = \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right) \Rightarrow \psi_X(t) = \mu t + \frac{\sigma^2 t^2}{2}$

$$d\psi_X(t)/dt|_{t=0} = [\mu + \sigma^2 t]|_{t=0} = \mu$$

$$d^2\psi_X(t)/dt^2|_{t=0} = \sigma^2|_{t=0} = \sigma^2$$

$$d^r\psi_X(t)/dt^r|_{t=0} = 0; r \geq 3$$

3–27. From Table XV, using the first column with scaling: $u_1 = 0.10480$, $u_2 = 0.22368$, $u_3 = 0.24130$, $u_4 = 0.42167$, $u_5 = 0.37570$, $u_6 = 0.77921, \dots, u_{20} = 0.07056$.

$$\begin{aligned}F_X(x) &= 0; \quad x < 0 \\ &= 0.5; \quad 0 \leq x < 1 \\ &= 0.75; \quad 1 \leq x < 2 \\ &= 0.875; \quad 2 \leq x < 3 \\ &= 1; \quad x \geq 3\end{aligned}$$

So, $x_1 = 0, x_2 = 0, x_3 = 0, x_4 = 0, x_5 = 0, x_6 = 2, \dots, x_{20} = 0$

3–28. The c.d.f. is

$$F_T(t) = 1 - e^{-t/4}$$

Thus, we set

$$t = -4\ln(1 - F_T(t)) = -4\ln(1 - u)$$

From Table XV, using the second column with scaling: $u_1 = 0.15011$, $u_2 = 0.46573$, $u_3 = 0.48360, \dots, u_{10} = 0.36257$

So,

$$\begin{aligned}t_1 &= -4\ell n(0.84989) = 0.6506 \\t_2 &= -4\ell n(0.53427) = 2.5074 \\\vdots \\t_{10} &= -4\ell n(0.63143) = 1.8391\end{aligned}$$