

14-1 It suffices to show that [see (14-41)]

$$H(A \cdot B | B_j) = H(A | B_j)$$

Since

$$A_i B_k B_j = \begin{cases} A_i B_j & k = j \\ \{\emptyset\} & k \neq j \end{cases} \quad \text{and } P(A_i B_j | B_j) = P(A_i | B_j)$$

(14-40) yields

$$\begin{aligned} H(A \cdot B | B_j) &= - \sum_{i,k} P(A_i B_k | B_j) \log P(A_i B_k | B_j) \\ &= - \sum_i P(A_i | B_j) \log P(A_i | B_j) = H(A | B_j) \end{aligned}$$

14-2 If $\alpha < \beta$, then $\phi'(\alpha) > \phi'(\beta)$ because

$$\phi'(\alpha) - \phi'(\beta) = \log(\beta/\alpha) > 0. \quad \text{Hence,}$$

$$\int_a^b \phi'(\alpha) d\alpha > \int_{a+c}^{b+c} \phi'(\alpha) d\alpha \quad c > 0$$

This yields

$$\phi(p_1 + p_2) - \phi(p_1) = \int_{p_1}^{p_1+p_2} \phi'(\alpha) d\alpha < \int_0^{p_2} \phi'(\alpha) d\alpha = \phi(p_2)$$

Similarly

$$\begin{aligned} &\phi(p_1 + \epsilon) - \phi(p_1) - \phi(p_2) + \phi(p_2 - \epsilon) \\ &= \int_{p_1}^{p_1+\epsilon} \phi'(\alpha) d\alpha - \int_{p_2-\epsilon}^{p_2} \phi'(\alpha) d\alpha > 0 \end{aligned}$$

14-3 Applying the identity

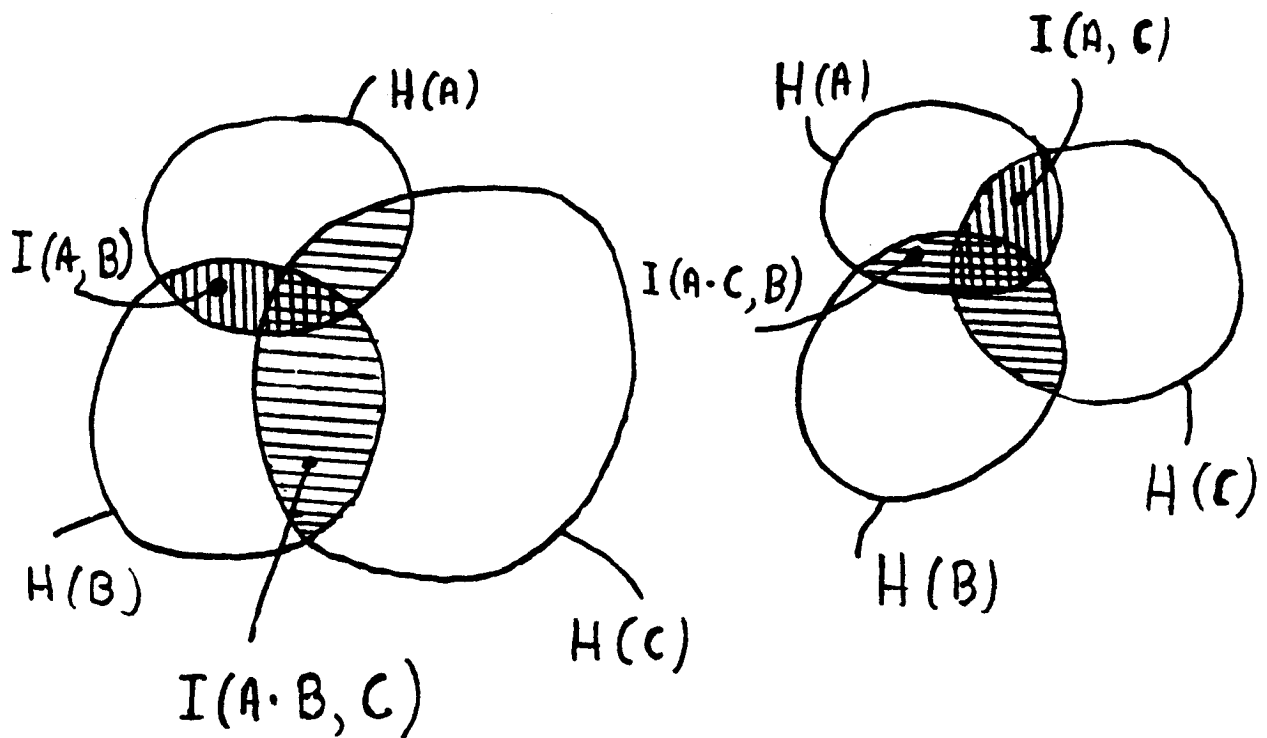
$$H(A_1 \cdot A_2) = H(A_1) + H(A_2|A_1) \quad (i)$$

to the partitions $A_1 = A$, $A_2 = B \cdot C$ and $A_1 = A \cdot B$, $A_2 = C$, we obtain the first line. The second line follows from the first [see (i)]. The third line is a consequence of the first two.

14-4 It follows if we apply the identity

$$I(A_1, A_2) = H(A_1) + H(A_2) - H(A_1 \cdot A_2)$$

to the partitions $A_1 = A \cdot B$, $A_2 = C$.



14-5 (a) From (14-53)

$$I(A, B \cdot C) = H(A) + H(B \cdot C) - H(A \cdot B \cdot C)$$

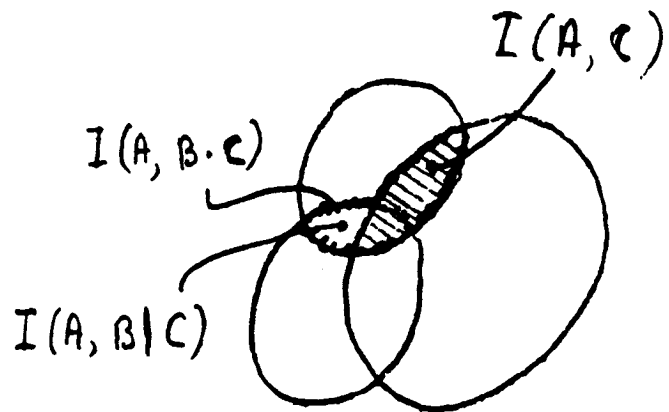
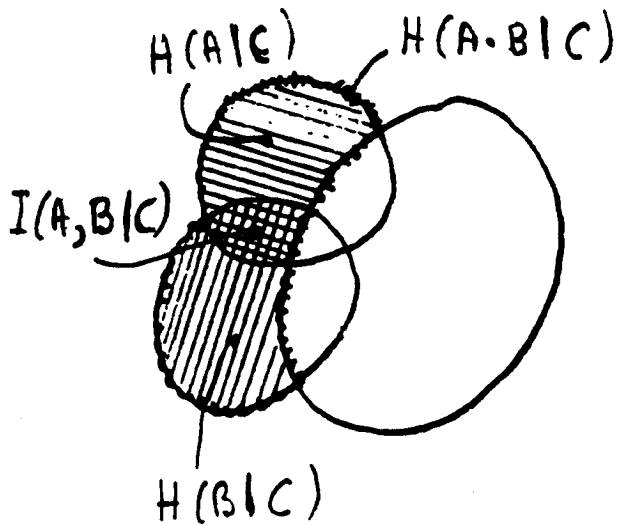
$$I(A, C) = H(A) + H(C) - H(A \cdot C)$$

and since (see Prob. 14-4)

$$H(A \cdot B \cdot C) - H(A \cdot C) = H(A \cdot B|C) - H(A|C)$$

we conclude with (14-49) that

$$I(A, B \cdot C) - I(A, C) = H(B|C) + H(A|C) - H(A \cdot B|C)$$



(b) If $B \cdot C$ is observed, then the resulting prediction in the uncertainty of A equals $I(A, B \cdot C)$. But, if $B \cdot C$ is observed, then C is observed, hence, the reduction in the uncertainty of A is at least $I(A, C)$. Hence

$$I(A, B \cdot C) \geq I(A, C)$$

with equality only if $I(A, B|C) = 0$, i.e., if in the subsequence of trials in which C occurred, knowledge of the occurrence of B gives no information about A .

14-6 The partition $H(A^3)$ has eight elements with respective probabilities

$$p^3, p^2q, p^2q, p^2q, pq^2, pq^2, pq^2, q^3$$

Hence

$$\begin{aligned} H(A^3) &= -p^3 \log p^3 - 3p^2q \log p^2q - 3pq^2 \log pq^2 - q^3 \log q^3 \\ &= -3p(p^2 + 2pq + q^2) \log p - 3q(p^2 + 2pq + q^2) \log q \\ &= -3p \log p - 3q \log q = 3H(A) \end{aligned}$$

14-7 The density of the RV $\underline{w} = \underline{x} + a$ equals $f_{\underline{x}}(w-a)$. Hence,

$$\begin{aligned} H(\underline{x} + a) &= - \int_{-\infty}^{\infty} f_{\underline{x}}(w-a) \log f_{\underline{x}}(w-a) dw \\ &= - \int_{-\infty}^{\infty} f_{\underline{x}}(x) \log f_{\underline{x}}(x) dx = H(\underline{x}) \end{aligned}$$

The joint density of the RVs \underline{x} and $\underline{z} = \underline{x} + \underline{y}$ equals $f_{\underline{xy}}(x, z-x)$. Hence [see (14-90)]

$$\begin{aligned} H(\underline{z} | \underline{x}) &= - \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{\underline{xy}}(x, z-x) \log f_{\underline{xy}}(x, z-x) / f_{\underline{x}}(x) dx dz \\ &= - \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{\underline{xy}}(x, y) \log f_{\underline{xy}}(x, y) / f_{\underline{x}}(x) dx dy = H(\underline{y} | \underline{x}) \end{aligned}$$

14-8 The RVs \underline{x} and \underline{y} take the values x_i and y_j respectively when $\underline{z} = x_i + y_j$ iff $\underline{x} = x_i$ and $\underline{y} = y_j$ (assumption). Hence,

$$\{\underline{z} = x_i + y_j\} = \{\underline{x} = x_i\} \cap \{\underline{y} = y_j\}$$

This shows that $A_z = A_x \cdot B_y$. Furthermore, since the RVs \underline{x} and \underline{y} are independent, the events $\{\underline{x} = x_i\}$ and $\{\underline{y} = y_j\}$ are also independent. This shows that the partitions A_x and B_y are independent and [see (14-44) and Prob. 14-1]

$$H(A_z | A_x) = H(A_x \cdot A_y | A_x) = H(A_y | A_x) = H(A_y)$$

From this it follows that $H(\underline{z} | \underline{x}) = H(\underline{y})$ because [see (14-88) and (14-41)]

$$H(\underline{z} | \underline{x}) = H(A_z | A_x)$$

14-9 As we see from (14-80)

$H(\underline{x}) = \ln a$ where we assume that $a = N\delta$. The RV \underline{y} takes the values $0, \delta, \dots, (N-1)\delta$ with probability $1/N$. The conditional density of \underline{x} assuming $\underline{y} = k\delta$ is uniform in the interval $(k\delta, k\delta + \delta)$. Hence,

$$H(\underline{x} | \underline{y} = k\delta) = - \int_{k\delta}^{k\delta + \delta} f(\underline{x} | \underline{y} = k\delta) \ln f(\underline{x} | \underline{y} = k\delta) dx = \ln \delta$$

And as in (14-41)

$$H(\underline{x} | \underline{y}) = \sum_{k=0}^N H(\underline{x} | \underline{y} = k\delta) P\{\underline{y} = k\delta\} = \ln \delta$$

Finally [see (14-95)]

$$I(\underline{x}, \underline{y}) = H(\underline{x}) - H(\underline{x} | \underline{y}) = \ln a - \ln \delta$$

14-10 If $y_i = g(x_i)$, $y_j = g(x_j)$ and $y_i = y_j$ then $x_i = x_j$. Hence,

$$p_{ij} = \begin{cases} p_i & i = j \\ 0 & i \neq j \end{cases} \quad p_i = P\{\underline{x} = x_i\}$$

and

$$H(\underline{x}, \underline{y}) = - \sum_{i,j} p_{ij} \log p_{ij} = - \sum_i p_i \log p_i = H(\underline{x})$$

14-11 From Prob. 10-10 it follows with $g(x) = x$ that $H(\underline{x}, \underline{x}) = H(\underline{x})$.
 And since [see (14-103)] $H(\underline{x}, \underline{x}) = H(\underline{x}|\underline{x}) + H(\underline{x})$ we conclude that
 $H(\underline{x}|\underline{x}) = 0$. From Prob. 14-3 it follows that

$$\begin{aligned} H(\underline{y}, \underline{x}|\underline{x}) &= H(A_y \cdot A_x | A_x) = H(A_x \cdot A_x) + H(A_y | A_x \cdot A_x) \\ &= H(A_y | A_x) = H(\underline{y}|\underline{x}) \end{aligned}$$

because $A_x \cdot A_x = A_x$ and $H(A_x \cdot A_x) = H(\underline{x}, \underline{x}) = 0$.

14-12 $E\{x_{-n}\} = 0$ $E\{x_{-n}^2\} = 5$ $E\{y_{-n}\} = 0$

$$E\{y_{-n}^2\} = \sum_{k=0}^{\infty} 2^{-2k} E\{x_{-n-k}^2\} = \frac{20}{3} \quad E\{x_{-n} y_{-n}\} = E\{x_{-n}^2\} = 5$$

(a) From (14-95), (14-84), and (15-86) with $\mu_{11} = 5$, $\mu_{22} = 20/3$,
 and $\mu_{12} = 5$

$$H(\underline{x}) = \ln \sqrt{10\pi e} \quad H(\underline{y}) = \ln \sqrt{40\pi e/3} \quad H(\underline{x}, \underline{y}) = \ln 10\pi e / \sqrt{3}$$

$$I(\underline{x}, \underline{y}) = \ln 2$$

(b) The process $y(t)$ is the output of the system

$$L(z) = \frac{1}{1 - 0.5 z^{-1}} \quad \ell_0 = 1$$

with input x_n . Since $\bar{H}(\underline{x}) = H(\underline{x})$ and [see (12A-1)]

$$\frac{1}{2\pi} \int_{-\pi}^{\pi} \ln |L(e^{j\phi})| d\phi = \ln \ell_0 = 0$$

(14-133) yields $\bar{H}(\underline{y}) = \bar{H}(\underline{x}) = H(\underline{x}) = \ln \sqrt{10\pi e}$.

14-13

$$\bar{H}(x) = H(x) = -\frac{1}{2} \int_4^6 \ln \frac{1}{2} dx = \ln 2$$

And as in Prob. 14-12 with $\ell_0 = 5$,

$$\bar{H}(y) = \bar{H}(x) + \ln 5 = \ln 10$$

14-14 Given that $f(x) = 0$ for $|x| > 1$ and $E(x) = 0.3$, find $f(x)$. With $g(x) = x$, (14-143) yields $f(x) = Ae^{-\lambda x}$ where

$$A \int_{-1}^1 e^{-\lambda x} dx = \frac{A}{\lambda} (e^\lambda - e^{-\lambda}) = 1$$

$$A \int_{-1}^1 xe^{-\lambda x} dx = \frac{A}{\lambda^2} (e^\lambda - e^{-\lambda}) - \frac{A}{\lambda} (e^\lambda - e^{-\lambda}) = 0.31$$

Solving, we obtain $A \approx 0.425$, $\lambda \approx -1$

14-15 $f(x) = Ae^{-\lambda x}$ for $1 < x < 5$ and 0 otherwise,

$$A \int_1^5 e^{-\lambda x} dx = 0.31 \quad A \int_1^5 xe^{-\lambda x} dx = 3 \frac{37}{60}$$

Hence, $A \approx 1.06$, $\lambda \approx 0.5$

14-16 From (14-151) with $x_k=k$, $g_1(x_k) = g_1(k) = k$, $k=1, \dots, 6$

$$g_2(x_k) = \begin{cases} 0 & k=1,3,5 \\ 1 & k=2,4,6 \end{cases} \quad P_k = \begin{cases} Ae^{-\lambda_1 k} & k=1,3,5 \\ Ae^{-\lambda_1 x - \lambda_2} & k=2,4,6 \end{cases}$$

Since $p_1 + p_3 + p_5 = 0.5$ and $E\{\underline{x}\} = 4.44$, we conclude with $z = e^{-\lambda_2}$ and $w = e^{-\lambda_1}$ that

$$A(z+z^3+z^5) = Aw(z^2+z^4+z^6)$$

$$A(z+3z^3+5z^5) + Aw(2z^2+4z^4+6z^6) = 4.44$$

This yields $A \simeq 0.0437$, $z = 1/w \simeq 1.468$

14-17 (a) The transformation $\underline{y} = 3\underline{x}$ is one-to-one, hence, $H(\underline{y}) = H(\underline{x})$

(b) From (14-113) with $g(x) = 3x$: $H(\underline{y}) = H(\underline{x}) + \ell n 3$

14-18 (a) For fair dice, $P\{7\} = \frac{1}{6}$, $P\{11\} = \frac{1}{18}$, $P\{\text{neither } 7 \text{ nor } 11\} = \frac{14}{18}$

$$H(A) = - \left[\frac{1}{6} \ell n \frac{1}{6} + \frac{1}{18} \ell n \frac{1}{18} + \frac{14}{18} \ell n \frac{14}{18} \right] = 0.655$$

(b) From (14-10) with $n=100$ and $N=3$:

$$n_T \simeq e^{nH(A)} \simeq 2.79 \times 10^{28} \quad n_a \simeq N^n \simeq 5.16 \times 10^{47}$$

14-19 The process \underline{x}_n is WSS with entropy rate $\bar{H}(x)$. Show that, if

$$\underline{w}_n = \sum_{k=0}^n \underline{x}_{n-k} \ell_k$$

then

$$\lim_{n \rightarrow \infty} \frac{1}{n+1} H(\underline{w}_0, \dots, \underline{w}_n) = \bar{H}(x) + \ln |\ell_0| \quad (i)$$

Proof. The RVs $\underline{w}_0, \dots, \underline{w}_n$ are linear transformations of the RVs $\underline{x}_0, \dots, \underline{x}_n$ and the transformation matrix equals

$$\begin{bmatrix} \ell_0 & 0 & \dots & 0 \\ \ell_1 & \ell_0 & \dots & 0 \\ \dots & \dots & \dots & \dots \\ \ell_n & \ell_{n-1} & \dots & 0 \end{bmatrix}$$

Since the determinant of this transformation equals $|\ell_0|^{n+1}$, (14-115) yields

$$H(\underline{w}_0, \dots, \underline{w}_n) = H(\underline{x}_0, \dots, \underline{x}_n) + (n+1) \ln |\ell_0|$$

Dividing by $(n+1)$ we obtain (i) as $n \rightarrow \infty$.

14-20 As in Example 14-19, $f(p) = A e^{-\lambda p}$. To find λ , we use the λ - η curve of Fig. 14-16. This yields

$$\lambda = -1.23 \quad f(p) = 0.51 e^{1.23p}$$

14-21 As in Example 14-22, $p_k = A e^{-\lambda k}$. To find λ , we use the w - η curve of Fig. 14-17. This yields (see also Jaynes)

$$w = 1.449 \quad \lambda = -0.371$$

p_1	p_2	p_3	p_4	p_5	p_6
0.054	0.079	0.114	0.165	0.240	0.348

14-22 The unknown density is normal as in (14-157) where

$$\Delta = \begin{vmatrix} 4 & 1 & 1 \\ 1 & 4 & m_{23} \\ 1 & m_{23} & 4 \end{vmatrix} = -4m_{23}^2 + 2m_{23} + 56$$

The moment $m_{23} = E\{x_2 x_3\}$ must be such as to maximize Δ . This yields $m_{23} = 0.25$.

14-23

Shannon

$$L = 2.7$$

p_i	0.3	0.2	0.15	0.15	0.1	0.06	0.04	
	$\frac{1}{4} \leq p_i < \frac{1}{2}$	$\frac{1}{8} \leq p_i < \frac{1}{4}$	$\frac{1}{16} \leq p_i < \frac{1}{8}$	$\frac{1}{32} \leq p_i < \frac{1}{16}$	$\frac{1}{64} \leq p_i < \frac{1}{32}$	$\frac{1}{128} \leq p_i < \frac{1}{64}$	$\frac{1}{256} \leq p_i < \frac{1}{128}$	$\sum_{i=1}^7 \frac{1}{2^{m_i}}$
m_i	2	3	3	3	4	5	5	0.75
	2	3	3	3	4	4	4	0.8125
	2	3	3	3	3	4	4	0.875
	2	3	3	3	3	3	4	0.9375
	2	3	3	3	3	3	3	1
x_i	00	010	011	100	101	110	111	

Fano
 $L = 2.6$

P_i	0.3	0.2	0.15	0.15	0.1	0.06	0.04
	A_0 0.5		A_2 0.5				
	A_{00} 0.3	A_{01} 0.2	A_{10} 0.3		A_{11} 0.2		
			A_{100} 0.15	A_{101} 0.15	A_{110} 0.1	A_{111} 0.1	
						A_{1110}	A_{1111}
x_i	00	01	100	101	110	1110	1111

Huffman
 $L = 2.6$

	1	2	3	4	5	6	7
	1	2	3	4	5	6	7
						0	1
	1	2	5	6	7		
			0	10	11	3	4
	1	3	4	2	5	6	7
		0	1		0	10	11
	2	5	6	7	1	3	4
	0	10	110	111		0	1
	1	3	4	2	5	6	7
	0	10	11	0	10	110	111
	1	3	4	2	5	6	7
	00	010	011	10	110	1110	1111
x_i	00	10	010	011	110	1110	1111

14-24 If $\underline{x}_n = 0$, then $\bar{x}_n = 000$ and $y_n = 1$ iff \bar{y}_n consists of one 0 or no zeros. The probability of one and only one zero equals $3\beta^2(1-\beta)$ [see (3-13)]; the probability of no zeros equals β^3 . Hence,

$$P\{y_n = 1 | \underline{x}_n = 0\} = 3\beta^2(1-\beta) + \beta^3$$

Thus, the redundantly coded channel of Example 14-29 is symmetrical as in (14-191) with probability of error $\beta_1 = \beta^2$.

14-25 If the received information is always wrong, then

$$P\{y_n = 1 | \underline{x}_n = 0\} = \beta = 1, \text{ hence } C = 1 - r(\beta) = 1$$
